



Genome-wide association study of more than 40,000 bipolar disorder cases provides new insights into the underlying biology

Bipolar disorder is a heritable mental illness with complex etiology. We performed a genome-wide association study of 41,917 bipolar disorder cases and 371,549 controls of European ancestry, which identified 64 associated genomic loci. Bipolar disorder risk alleles were enriched in genes in synaptic signaling pathways and brain-expressed genes, particularly those with high specificity of expression in neurons of the prefrontal cortex and hippocampus. Significant signal enrichment was found in genes encoding targets of antipsychotics, calcium channel blockers, antiepileptics and anesthetics. Integrating expression quantitative trait locus data implicated 15 genes robustly linked to bipolar disorder via gene expression, encoding druggable targets such as HTR6, MCHR1, DCLK3 and FURIN. Analyses of bipolar disorder subtypes indicated high but imperfect genetic correlation between bipolar disorder type I and II and identified additional associated loci. Together, these results advance our understanding of the biological etiology of bipolar disorder, identify novel therapeutic leads and prioritize genes for functional follow-up studies.

Bipolar disorder (BD) is a complex mental disorder characterized by recurrent episodes of (hypo)mania and depression. It is a common condition affecting an estimated 40 to 50 million people worldwide¹. This, combined with the typical onset in young adulthood, an often chronic course and increased risk of suicide², makes BD a major public health concern and a major cause of global disability¹. Clinically, BD is classified into two main subtypes: bipolar I disorder (BD I), in which manic episodes typically alternate with depressive episodes, and bipolar II disorder (BD II), characterized by the occurrence of at least one hypomanic and one depressive episode³. These subtypes have a lifetime prevalence of ~1% each in the population^{4,5}.

Family and molecular genetic studies provide convincing evidence that BD is a multifactorial disorder, with genetic and environmental factors contributing to its development⁶. On the basis of twin and family studies, the heritability of BD is estimated at 60–85%^{7,8}. Genome-wide association studies (GWAS)^{9–23} have led to valuable insights into the genetic etiology of BD. The largest such study has been conducted by the Psychiatric Genomics Consortium (PGC), in which genome-wide SNP data from 29,764 patients with BD and 169,118 controls were analyzed and 30 genome-wide significant loci were identified (PGC2)²⁴. SNP-based heritability (h^2_{SNP}) estimation using the same data suggested that common genetic variants genome-wide explain ~20% of BD's phenotypic variance²⁴. Polygenic risk scores (PRSs) generated from the results of this study explained ~4% of phenotypic variance in independent samples. Across the genome, genetic associations with BD converged on specific biological pathways including regulation of insulin secretion^{25,26}, retrograde endocannabinoid signaling²⁴, glutamate receptor signaling²⁷ and calcium channel activity⁹.

Despite this considerable progress, only a fraction of the genetic etiology of BD has been identified, and the specific biological mechanisms underlying the development of the disorder are still unknown. In the present study, we report the results of the third GWAS meta-analysis of the PGC Bipolar Disorder Working Group, comprising 41,917 individuals with BD and 371,549 controls. These results confirm and expand on many previously reported findings, identify novel therapeutic leads and prioritize genes for functional

follow-up studies^{28,29}. Thus, our results further illuminate the biological etiology of BD.

Results

GWAS results. We conducted a GWAS meta-analysis of 57 BD cohorts collected in Europe, North America and Australia (Supplementary Table 1), totaling 41,917 individuals with BD (cases) and 371,549 controls of European descent (effective $n=101,962$, see Methods). For 52 cohorts, individual-level genotype and phenotype data were shared with the PGC and cases met international consensus criteria (DSM-IV, ICD-9 or ICD-10) for lifetime BD, established using structured diagnostic interviews, clinician-administered checklists or medical record review. BD GWAS summary statistics were received for five external cohorts (iPSYCH³⁰, deCODE genetics³¹, Estonian Biobank³², Trøndelag Health Study (HUNT)³³ and UK Biobank³⁴), in which most cases were ascertained using ICD codes. The GWAS meta-analysis identified 64 independent loci associated with BD at genome-wide significance ($P < 5 \times 10^{-8}$; Fig. 1, Table 1 and Supplementary Table 2). Using linkage disequilibrium score regression (LDSC)³⁵, the h^2_{SNP} of BD was estimated to be 18.6% (s.e.=0.008, $P=5.1 \times 10^{-132}$) on the liability scale, assuming a BD population prevalence of 2%, and 15.6% (s.e.=0.006, $P=5.0 \times 10^{-132}$) assuming a population prevalence of 1% (Supplementary Table 3). The genomic inflation factor (λ_{GC}) was 1.38 and the LDSC intercept was 1.04 (s.e.=0.01, $P=2.5 \times 10^{-4}$; Supplementary Fig. 1). While the intercept has frequently been used as an indicator of confounding from population stratification, it can rise above 1 with increased sample size and heritability. The attenuation ratio—(LDSC intercept – 1)/(mean of association chi-square statistics – 1)—which is not subject to these limitations, was 0.06 (s.e.=0.02), indicating that the majority of inflation of the GWAS test statistics was due to polygenicity^{35,36}. Of the 64 genome-wide significant loci, 33 are novel discoveries (that is, loci not overlapping with any locus previously reported as genome-wide significant for BD). Novel loci include the major histocompatibility complex (MHC) and loci previously reaching genome-wide significance for other psychiatric disorders, including ten for schizophrenia, four for major depression and three for childhood-onset psychiatric disorders or problematic alcohol use (Table 1).

A full list of authors and their affiliations appears at the end of the paper.

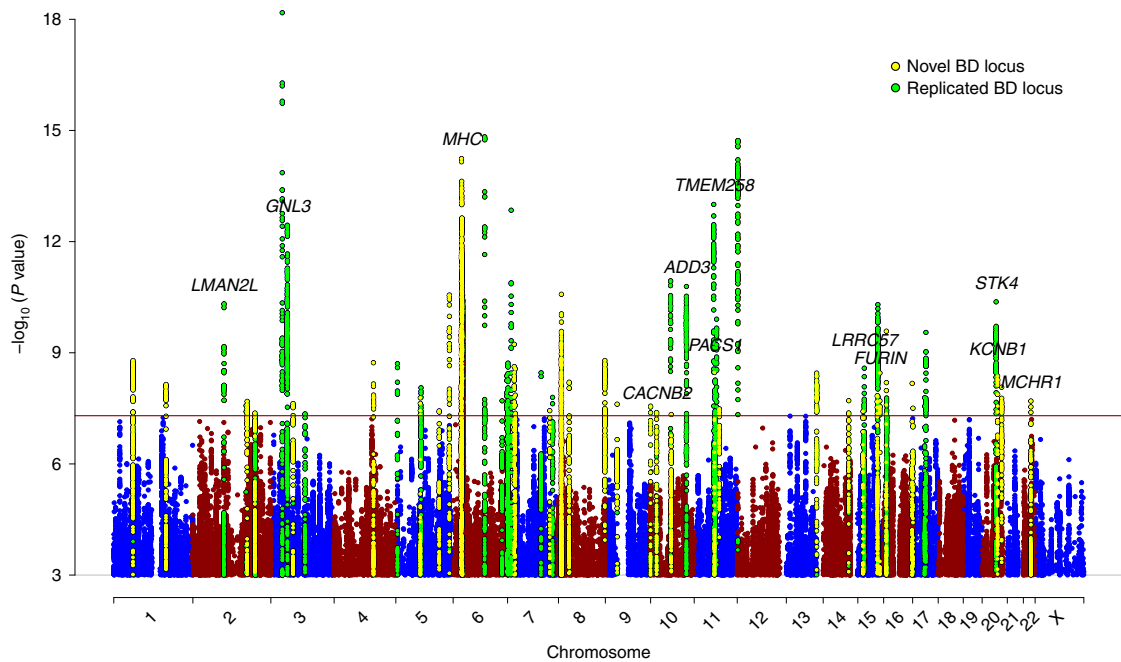


Fig. 1 | Manhattan plot of genome-wide association meta-analysis of 41,917 BD cases and 371,549 controls. The x axis shows genomic position (chromosomes 1–22 and X), and the y axis shows statistical significance as $-\log_{10}[P \text{ value}]$. P values are two-sided and based on an inverse-variance-weighted fixed-effects meta-analysis. The red line shows the genome-wide significance threshold ($P < 5 \times 10^{-8}$). SNPs in genome-wide significant loci are colored green for loci previously associated with BD and yellow for novel associations from this study. The genes labeled are those prioritized by integrative eQTL analyses or notable genes in novel loci (MHC, CACNB2 and KCNB1).

Enrichment analyses. Genome-wide analyses using MAGMA³⁷ indicated significant enrichment of BD associations in 161 genes (Supplementary Table 4) and 4 gene sets related to synaptic signaling (Supplementary Table 5). The BD association signal was enriched among genes expressed in different brain tissues (Supplementary Table 6), especially genes with high specificity of gene expression in neurons (both excitatory and inhibitory) versus other cell types, within cortical and subcortical brain regions in mice (Supplementary Fig. 2)³⁸. In human brain samples, signal enrichment was also observed in hippocampal pyramidal neurons and interneurons of the prefrontal cortex and hippocampus, compared with other cell types (Supplementary Fig. 2).

In a gene-set analysis of the targets of individual drugs (from the Drug–Gene Interaction Database DGIdb v.2 (ref.³⁹) and the Psychoactive Drug Screening Database Ki DB⁴⁰), the targets of the calcium channel blockers mibefradil and nisoldipine were significantly enriched (Supplementary Table 7). Grouping drugs according to their anatomical therapeutic chemical classes⁴¹, there was significant enrichment in the targets of four broad drug classes (Supplementary Table 8): psycholeptics (drugs with a calming effect on behavior; especially hypnotics and sedatives, antipsychotics and anxiolytics), calcium channel blockers, antiepileptics and (general) anesthetics (Supplementary Table 8).

Expression quantitative trait locus integrative analyses. We conducted a transcriptome-wide association study (TWAS) using FUSION⁴² and expression quantitative trait locus (eQTL) data from the PsychENCODE Consortium (1,321 brain samples)⁴³. BD-associated alleles significantly influenced expression of 77 genes in the brain (Supplementary Table 9 and Supplementary Fig. 3). These genes encompassed 40 distinct regions. We performed TWAS fine-mapping using FOCUS⁴⁴ to model the correlation among the TWAS signals and prioritize the most likely causal gene(s) in each region. Within the 90%-credible set, FOCUS prioritized 22 genes with a posterior inclusion probability (PIP) > 0.9 (encompassing 20

distinct regions) and 32 genes with a PIP > 0.7 (29 distinct regions; Supplementary Table 10).

We used summary-data-based Mendelian randomization (SMR)^{45,46} to identify putative causal relationships between SNPs and BD via gene expression by integrating the BD GWAS results with brain eQTL summary statistics from the PsychENCODE⁴³ Consortium and blood eQTL summary statistics from the eQTLGen consortium (31,684 whole blood samples)⁴⁷. The eQTLGen results represent the largest existing eQTL study and provide independent eQTL data. Of the 32 genes fine-mapped with PIP > 0.7 , 15 were significantly associated with BD in the SMR analyses and passed the heterogeneity in dependent instruments (HEIDI) test^{45,46}, suggesting that their effect on BD is mediated via gene expression in the brain and/or blood (Supplementary Table 11). The genes located in genome-wide significant loci are labeled in Fig. 1. Other significant genes included *HTR6*, *DCLK3*, *HAPLN4* and *PACSIN2*.

MHC locus. Variants within and distal to the MHC locus were associated with BD at genome-wide significance. The most highly associated SNP was *rs13195402*, 3.2 megabases (Mb) distal to any *HLA* gene or the complement component 4 (*C4*) genes (Supplementary Fig. 4). Imputation of *C4* alleles using SNP data uncovered no association between the five most common structural forms of the *C4A/C4B* locus (BS, AL, AL–BS, AL–BL and AL–AL) and BD, either before or after conditioning on *rs13195402* (Supplementary Fig. 5). While genetically predicted *C4A* expression initially showed a weak association with BD, this association was nonsignificant after controlling for *rs13195402* (Supplementary Fig. 6).

Polygenic risk scoring. The performance of PRSs based on these GWAS results was assessed by excluding cohorts in turn from the meta-analysis to create independent test samples. PRSs explained ~4.57% of phenotypic variance in BD on the liability scale (at GWAS P -value threshold (GWAS P_T) < 0.1 , BD population prevalence 2%), based on the weighted mean R^2 across cohorts (Fig. 2

Table 1 | Genome-wide significant loci for BD from meta-analysis of 41,917 cases and 371,549 controls

Locus	CHR	BP	SNP	P	OR	s.e.	A1/A2	A1 freq in controls	Previous report ^a for BD (citation)	Name for novel locus ^b	Previous report ^a for psychiatric disorders
1	1	61105668	rs2126180	1.6 × 10 ⁻⁹	1.058	0.009	A/G	0.457		LINC01748	
2	1	163745389	rs10737496	7.2 × 10 ⁻⁹	1.056	0.009	C/T	0.444		NUF2	CDG
3 ^c	2	97416153	rs4619651	4.8 × 10 ⁻¹¹	1.068	0.010	G/A	0.670	LMAN2L (PGC2)		CDG
4	2	166152389	rs17183814	2.7 × 10 ⁻⁸	1.108	0.019	G/A	0.924	SCN2A (PGC2)		
5	2	169481837	rs13417268	2.1 × 10 ⁻⁸	1.064	0.011	C/G	0.758		CERS6	
6	2	193738336	rs2011302	4.3 × 10 ⁻⁸	1.055	0.010	A/T	0.377		PCGEM1	CDG
7	2	194437889	rs2719164	4.9 × 10 ⁻⁸	1.053	0.010	A/G	0.564	Intergenic (PGC2)		CDG
8 ^c	3	36856030	rs9834970	6.6 × 10 ⁻¹⁹	1.087	0.009	C/T	0.481	TRANK1 (PGC2)		SCZ, CDG
9 ^c	3	52626443	rs2336147	3.6 × 10 ⁻¹³	1.070	0.009	T/C	0.498	ITIH1 (PGC2)		SCZ, CDG
10	3	70488788	rs115694474	2.4 × 10 ⁻⁸	1.068	0.012	T/A	0.799		MDFIC2	
11	3	107757060	rs696366	4.5 × 10 ⁻⁸	1.053	0.009	C/A	0.550	CD47 (PGC2)		
12 ^c	4	123076007	rs112481526	1.9 × 10 ⁻⁹	1.065	0.011	G/A	0.256		KIAA1109	MD
13 ^c	5	7542911	rs28565152	2.0 × 10 ⁻⁹	1.070	0.011	A/G	0.238	ADCY2 (PGC2)		
14 ^c	5	78849505	rs6865469	1.7 × 10 ⁻⁸	1.060	0.010	T/G	0.274		HOMER1	
15	5	80961069	rs6887473	8.8 × 10 ⁻⁹	1.062	0.011	G/A	0.739	SSBP2 (PGC2)		
16 ^c	5	137712121	rs10043984	3.7 × 10 ⁻⁸	1.062	0.011	T/C	0.236		KDM3B	CDG
17	5	169289206	rs10866641	2.8 × 10 ⁻¹¹	1.065	0.009	T/C	0.575		DOCK2	
18 ^c	6	26463575	rs13195402	5.8 × 10 ⁻¹⁵	1.146	0.018	G/T	0.919		MHC	MD, SCZ, CDG, MOOD
19 ^c	6	98565211	rs1487445	1.5 × 10 ⁻¹⁵	1.078	0.009	T/C	0.487	POU3F2 (PGC2)		CDG
20	6	152793572	rs4331993	2.0 × 10 ⁻⁸	1.056	0.010	A/T	0.382	SYNE1 (Green et al. 2013)		
21 ^c	6	166995260	rs10455979	4.2 × 10 ⁻⁹	1.057	0.010	G/C	0.500	RPS6KA2 (PGC2)		
22 ^c	7	2020995	rs12668848	1.9 × 10 ⁻⁹	1.059	0.010	G/A	0.575	MAD1L1 (Hou et al. 2016, Ikeda et al. 2018)		MD, SCZ, CDG
23 ^c	7	11871787	rs113779084	1.4 × 10 ⁻¹³	1.079	0.010	A/G	0.299	THSD7A (PGC2)		
24 ^c	7	21492589	rs6954854	5.9 × 10 ⁻¹⁰	1.060	0.009	G/A	0.425		SP4	
25	7	24647222	rs12672003	2.7 × 10 ⁻⁹	1.096	0.016	G/A	0.113		MPP6	SCZ, CDG, MOOD
26	7	105043229	rs11764361	3.5 × 10 ⁻⁹	1.063	0.010	A/G	0.668	SRPK2 (PGC2)		SCZ, ASD, CDG
27	7	131870597	rs6946056	3.7 × 10 ⁻⁸	1.055	0.010	C/A	0.623		PLXNA4	
28	7	140676153	rs10255167	1.6 × 10 ⁻⁸	1.068	0.012	A/G	0.778	MRPS33 (PGC2)		CDG
29 ^c	8	9763581	rs62489493	2.6 × 10 ⁻¹¹	1.094	0.014	G/C	0.128		miR124-1	SCZ, ALC, ASD
30 ^c	8	10226355	rs3088186	2.1 × 10 ⁻⁸	1.058	0.010	T/C	0.287		MSRA	SCZ, ALC, ASD
31	8	34152492	rs2953928	6.3 × 10 ⁻⁹	1.124	0.020	A/G	0.067		RP1-84O15.2 (lincRNA)	SCZ, ADHD, CDG
32 ^c	8	144993377	rs6992333	1.6 × 10 ⁻⁹	1.062	0.010	G/A	0.410		PLEC	
33	9	37090538	rs10973201	2.5 × 10 ⁻⁸	1.101	0.017	C/T	0.110		ZCCHC7	MD, CDG, MOOD
34 ^c	9	141066490	rs62581014	2.8 × 10 ⁻⁸	1.067	0.012	T/C	0.366		TUBBP5	
35 ^c	10	18751103	rs1998820	4.1 × 10 ⁻⁸	1.087	0.015	T/A	0.886		CACNB2	SCZ, CDG
36 ^c	10	62322034	rs10994415	1.1 × 10 ⁻¹¹	1.125	0.017	C/T	0.082	ANK3 (PGC2)		
37	10	64525135	rs10761661	4.7 × 10 ⁻⁸	1.053	0.009	T/C	0.472		ADO	

Continued

Table 1 | Genome-wide significant loci for BD from meta-analysis of 41,917 cases and 371,549 controls (Continued)

Locus	CHR	BP	SNP	P	OR	s.e.	A1/ A2	A1 freq in controls	Previous report ^a for BD (citation)	Name for novel locus ^b	Previous report ^a for psychiatric disorders
38 ^c	10	111648659	rs2273738	1.6 × 10 ⁻¹¹	1.096	0.014	T/C	0.135	ADD3 (Charney et al. 2017, PGC2)		
39 ^c	11	61618608	rs174592	9.9 × 10 ⁻¹⁴	1.074	0.010	G/A	0.395	FADS2 (PGC2)		MD, CDG, MOOD
40	11	64009879	rs4672	3.4 × 10 ⁻⁹	1.107	0.017	A/G	0.083		FKBP2	
41 ^c	11	65848738	rs475805	2.0 × 10 ⁻⁹	1.070	0.011	A/G	0.767	PACS1 (PGC2)		
42 ^c	11	66324583	rs678397	5.5 × 10 ⁻⁹	1.056	0.009	T/C	0.457	PC (PGC1, PGC2)		
43 ^c	11	70517927	rs12575685	1.2 × 10 ⁻¹⁰	1.067	0.010	A/G	0.327	SHANK2 (PGC2)		MD
44	11	79092527	rs12289486	3.3 × 10 ⁻⁸	1.086	0.015	T/C	0.115	ODZ4 (PGC1)		
45 ^c	12	2348844	rs11062170	1.9 × 10 ⁻¹⁵	1.081	0.010	C/G	0.333	CACNA1C (PGC2)		SCZ, CDG, MOOD
46	13	113869045	rs35306827	3.6 × 10 ⁻⁹	1.068	0.011	G/A	0.775		CUL4A	
47	14	99719219	rs2693698	2.0 × 10 ⁻⁸	1.055	0.009	G/A	0.551		BCL11B	SCZ, CDG
48 ^c	15	38973793	rs35958438	3.8 × 10 ⁻⁸	1.066	0.012	G/A	0.772		C15orf53	CDG
49 ^c	15	42904904	rs4447398	2.6 × 10 ⁻⁹	1.086	0.014	A/C	0.131	STARD9 (PGC2)		
50	15	83531774	rs62011709	1.4 × 10 ⁻⁸	1.064	0.011	T/A	0.747		HOMER2	SCZ
51 ^c	15	85149575	rs748455	5.0 × 10 ⁻¹¹	1.070	0.010	T/C	0.719	ZNF592 (PGC2)		SCZ, CDG
52	15	91426560	rs4702	3.5 × 10 ⁻⁹	1.059	0.010	G/A	0.446		FURIN	SCZ, CDG
53	16	9230816	rs28455634	2.6 × 10 ⁻¹⁰	1.065	0.010	G/A	0.620		C16orf72	
54	16	9926348	rs7199910	1.7 × 10 ⁻⁸	1.057	0.010	G/T	0.312	GRIN2A (PGC2)		SCZ, CDG
55	16	89632725	rs12932628	6.7 × 10 ⁻⁹	1.058	0.010	T/G	0.487		RPL13	
56	17	1835482	rs4790841	3.1 × 10 ⁻⁸	1.075	0.013	T/C	0.151		RTN4RL1	
57	17	38129841	rs11870683	2.8 × 10 ⁻⁸	1.059	0.010	T/A	0.650	ERBB2 (Hou et al. 2016)		
58	17	38220432	rs61554907	1.6 × 10 ⁻⁸	1.091	0.015	T/G	0.124	ERBB2 (Hou et al. 2016)		
59 ^c	17	42191893	rs228768	2.8 × 10 ⁻¹⁰	1.067	0.010	G/T	0.294	HDAC5 (PGC2)		
60 ^c	20	43682551	rs67712855	4.2 × 10 ⁻¹¹	1.070	0.010	T/G	0.687	STK4 (PGC2)		
61 ^c	20	43944323	rs6032110	1.0 × 10 ⁻⁹	1.059	0.009	A/G	0.512	WFDC12 (PGC2)		
62 ^c	20	48033127	rs237460	4.3 × 10 ⁻⁹	1.057	0.009	T/C	0.412		KCNB1	CDG
63	20	60865815	rs13044225	8.5 × 10 ⁻⁹	1.056	0.010	G/A	0.440		OSBPL2	
64	22	41153879	rs5758064	2.0 × 10 ⁻⁸	1.054	0.009	T/C	0.523		SLC25A17	MD, SCZ, CDG, MOOD

^aPrevious report refers to previous association of a SNP in the locus with the psychiatric disorder at genome-wide significance. PGC1, ref. ³; PGC2, ref. ²⁴; Hou et al. 2016, ref. ¹¹; Ikeda et al. 2018, ref. ⁵⁹; Green et al. 2013, ref. ⁶; Charney et al. 2017, ref. ¹¹. ^bNovel loci are named using the nearest gene to the index SNP. ^cLocus overlaps with a genome-wide significant locus for BD I. CHR, chromosome; BP, GRCh37 base-pair position; OR, odds ratio; A1, tested allele; A2, other allele; freq, frequency; CDG, cross-disorder GWAS of the PGC; MD, major depression; SCZ, schizophrenia; MOOD, mood disorders; ASD, autism spectrum disorder; ALC, alcohol use disorder or problematic alcohol use; ADHD, attention deficit/hyperactivity disorder. P values are two-sided and based on an inverse-variance-weighted fixed-effects meta-analysis.

and Supplementary Table 12). This corresponds to a weighted mean area under the curve of 65%. Results per cohort and per wave of recruitment to the PGC are in Supplementary Tables 12 and 13 and Supplementary Fig. 7. At GWAS $P_T < 0.1$, individuals in the top 10% of BD PRSs had an odds ratio of 3.5 (95% confidence interval (CI) 1.7–7.3) of being affected with the disorder compared with individuals in the middle decile (based on the weighted mean OR across PGC cohorts), and an odds ratio of 9.3 (95% CI 1.7–49.3) compared with individuals in the lowest decile. The generalizability of PRSs from this meta-analysis was examined in several non-European cohorts. PRSs explained up to 2.3% and 1.9% of variance in BD in two East Asian samples, and 1.2% and 0.4% in two admixed African American samples (Fig. 2 and Supplementary Table 14).

The variance explained by the PRSs increased in every cohort with increasing sample size of the PGC BD European discovery sample (Supplementary Fig. 8 and Supplementary Table 14).

Genetic architecture of BD and other traits. The genome-wide genetic correlation (r_g) of BD with a range of diseases and traits was assessed on LD Hub⁴⁸. After correction for multiple testing, BD showed significant r_g with 16 traits among 255 tested from published GWASs (Supplementary Table 15). Genetic correlation was positive with all psychiatric disorders assessed, particularly schizophrenia ($r_g = 0.68$) and major depression ($r_g = 0.44$), and to a lesser degree anorexia, attention deficit/hyperactivity disorder and autism spectrum disorder ($r_g \approx 0.2$). We found evidence of positive r_g between BD

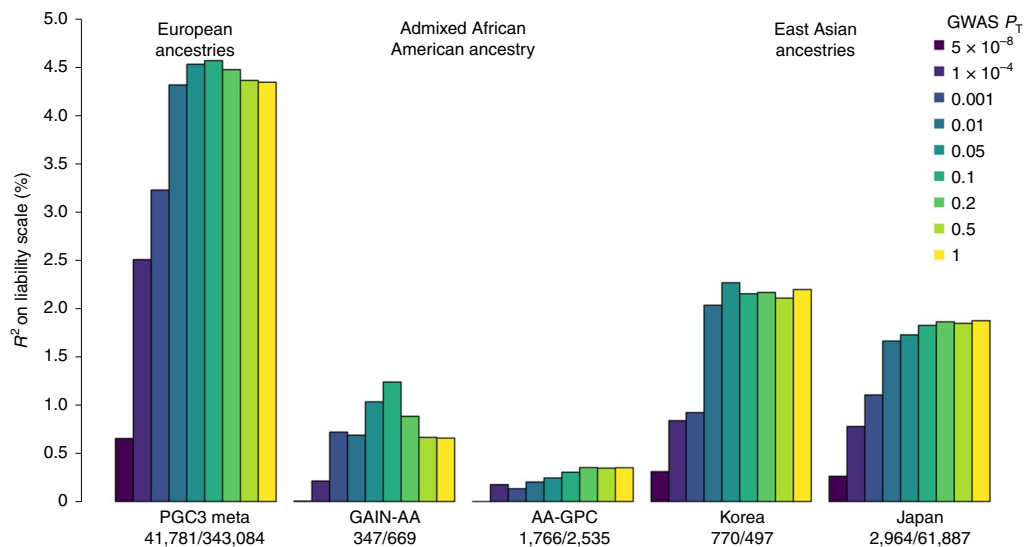


Fig. 2 | Phenotypic variance in BD explained by PRSs. Variance explained is presented on the liability scale, assuming a 2% population prevalence of BD. For European ancestries, the results shown are the weighted mean R^2 values across all 57 cohorts in the PGC3 meta-analysis, weighted by the effective n per cohort. The numbers of cases and controls are shown from left to right under the bar plot for each study. GWAS P_T , the color of the bars represents the P -value threshold used to select SNPs from the discovery GWAS; GAIN-AA, Genetic Association Information Network African American cohort; AA-GPC, African American Genomic Psychiatry Cohort.

and smoking initiation, cigarettes per day, problematic alcohol use and drinks per week (Fig. 3). BD was also positively genetically correlated with measures of sleep quality (daytime sleepiness, insomnia and sleep duration; Fig. 3). Among 514 traits measured in the general population of the UK Biobank, there was significant r_g between BD and many psychiatric-relevant traits or symptoms, dissatisfaction with interpersonal relationships, poorer overall health rating and feelings of loneliness or isolation (Supplementary Table 16).

Bivariate gaussian mixture models were applied to the GWAS summary statistics for BD and other complex traits using the MiXeR tool^{49,50} to estimate the number of variants influencing each trait that explain 90% of h^2_{SNP} and their overlap between traits. MiXeR estimated that approximately 8,600 (s.e. = 200) variants influence BD, which is similar to the estimate for schizophrenia (9,700, s.e. = 200) and lower than that for major depression (12,300, s.e. = 600; Supplementary Table 17 and Supplementary Fig. 9). When considering the number of shared loci as a proportion of the total polygenicity of each trait, the vast majority of loci influencing BD were also estimated to influence major depression (97%) and schizophrenia (96%; Supplementary Table 17 and Supplementary Fig. 9). Interestingly, within these shared components, the variants that influenced both BD and schizophrenia had high concordance in direction of effect (80%, s.e. = 2%), while the portion of concordant variants between BD and major depressive disorder was only 69% (s.e. = 1%; Supplementary Table 17).

Genetic and causal relationships between BD and modifiable risk factors. Ten traits associated with BD from clinical and epidemiological studies were investigated in detail for genetic and potentially causal relationships with BD via LDSC³⁵, generalized summary statistics-based Mendelian randomization (GSMR)⁵¹ and bivariate gaussian mixture modeling⁴⁹. BD has been strongly linked with sleep disturbances⁵², alcohol use⁵³, smoking⁵⁴, higher educational attainment^{55,56} and mood instability⁵⁷. Most of these traits had modest but significant genetic correlations with BD ($r_g = -0.05$ to 0.35 ; Fig. 3). Examining the effects of these traits on BD via GSMR, smoking initiation was associated with BD, corresponding to an OR of 1.49 (95% CI 1.38–1.61) for developing the disorder ($P = 1.74 \times 10^{-22}$;

Fig. 3). Testing the effect of BD on the traits, we found that BD was significantly associated with reduced likelihood of being a morning person and increased number of drinks per week ($P < 1.47 \times 10^{-3}$; Fig. 3). Positive bidirectional relationships were identified between BD and longer sleep duration, problematic alcohol use, educational attainment and mood instability (Fig. 3). Notably, the instrumental variables for mood instability were selected from a GWAS conducted in the general population, excluding individuals with psychiatric disorders⁵⁸. For all of the aforementioned BD–trait relationships, the effect size estimates from GSMR were consistent with those calculated using the inverse-variance-weighted regression method, and there was no evidence of bias from horizontal pleiotropy. Full MR results are in Supplementary Tables 18 and 19. Bivariate gaussian mixture modeling using MiXeR indicated large proportions of variants influencing both BD and all other traits tested, particularly educational attainment, where approximately 98% of variants influencing BD were estimated to also influence educational attainment. While cigarettes per day was a trait of interest, MiXeR could not model these data due to low polygenicity and heritability, and the effect of cigarettes per day on BD was inconsistent between MR methods, suggesting a violation of MR assumptions (Supplementary Tables 18–20).

BD subtypes. We conducted GWAS meta-analyses of BD I (25,060 cases, 449,978 controls) and BD II (6,781 cases, 364,075 controls). The BD I analysis identified 44 genome-wide significant loci, 31 of which overlapped with genome-wide significant loci from the main BD GWAS (Table 1 and Supplementary Table 21). The remaining 13 genome-wide significant loci for BD I all had $P < 4.0 \times 10^{-5}$ in the main BD GWAS. One genome-wide significant locus was identified in the GWAS meta-analysis of BD II and had a $P < 1.1 \times 10^{-4}$ in the main GWAS of BD (Supplementary Table 21). The h^2_{SNP} estimates on the liability scale for BD I and BD II were 20.9% (s.e. = 0.009, $P = 1.0 \times 10^{-111}$) and 11.6% (s.e. = 0.01, $P = 3.9 \times 10^{-15}$), respectively, assuming a 1% population prevalence of each subtype. These heritability values are significantly different from each other ($P = 2.4 \times 10^{-25}$, block jack-knife). The genetic correlation between BD I and BD II was 0.85

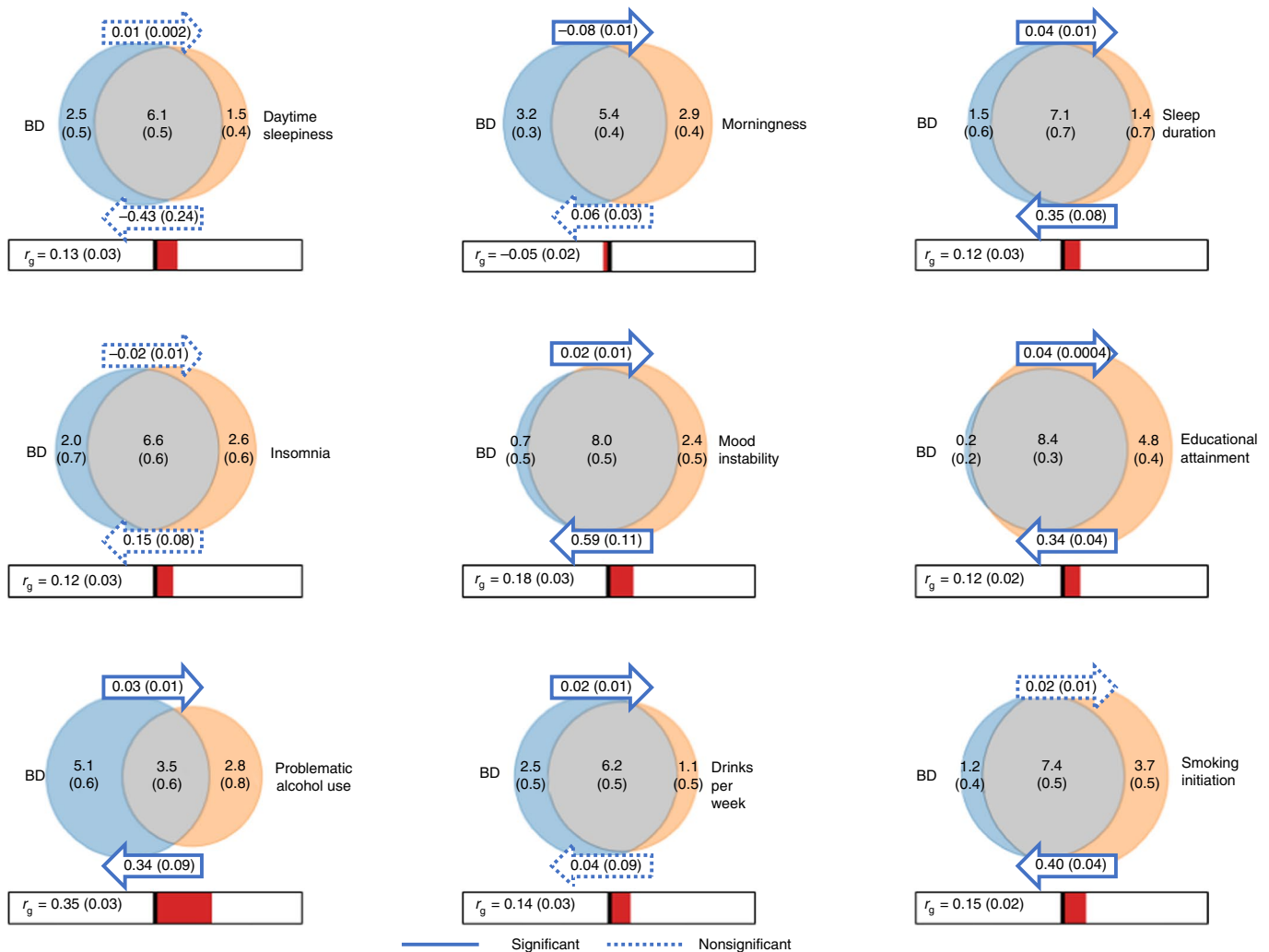


Fig. 3 | Relationships between BD and modifiable risk factors based on genetic correlations, GSMR and bivariate gaussian mixture modeling. The Venn diagrams depict MiXeR results of the estimated number of influencing variants shared between BD and each trait of interest (gray), unique to BD (blue) and unique to the trait of interest (orange). The number of influencing variants and standard error are shown in thousands. The size of the circles reflects the polygenicity of each trait, with larger circles corresponding to greater polygenicity. The estimated genetic correlation (r_g) between BD and each trait of interest and standard error from LDSC is shown below the corresponding Venn diagram, with an accompanying scale (–1 to +1). The arrows above and below the Venn diagrams indicate the results of GSMR of BD on the trait of interest, and the trait of interest on BD, respectively. The GSMR effect size and standard error is shown inside the corresponding arrow. Solid arrows indicate a significant relationship between the exposure and the outcome, after correction for multiple comparisons ($P < 1.47 \times 10^{-3}$), and dotted arrows indicate a nonsignificant relationship.

(s.e. = 0.05, $P = 2.88 \times 10^{-54}$), which is significantly different from 1 ($P = 1.6 \times 10^{-3}$). The genetic correlation of BD I with schizophrenia ($r_g = 0.66$, s.e. = 0.02) was higher than that of BD II ($r_g = 0.54$, s.e. = 0.05), whereas major depression was more strongly genetically correlated with BD II ($r_g = 0.66$, s.e. = 0.05) than with BD I ($r_g = 0.34$, s.e. = 0.03; Supplementary Table 22).

Discussion

In a GWAS of 41,917 BD cases, we identified 64 associated genomic loci, 33 of which are novel discoveries. With a 1.5-fold increase in effective sample size compared with the PGC2 BD GWAS, this study more than doubled the number of associated loci, representing an inflection point in the rate of risk variant discovery. We observed consistent replication of known BD loci, including 28/30 loci from the PGC2 GWAS²⁴ and several implicated by other BD GWAS^{15–17}, including a study of East Asian cases⁵⁹.

The 33 novel loci discovered here encompass genes of expected biological relevance to BD, such as the ion-channel-encoding

genes *CACNB2* and *KCNB1*. Among the 64 BD loci, 17 have previously been implicated in GWAS of schizophrenia⁶⁰, and 7 in GWAS of major depression⁶¹, representing the first overlap of genome-wide significant loci between the mood disorders. For these genome-wide significant loci shared across disorders, 17/17 and 5/7 of the BD index SNPs had the same direction of effect on schizophrenia and major depression, respectively (Supplementary Table 23). More generally, 50/64 and 62/64 BD loci had a consistent direction of effect on major depression and schizophrenia, respectively, considerably greater than chance ($P < 1 \times 10^{-5}$, binomial test). Bivariate gaussian mixture modeling estimated that across the entire genome, almost all variants influencing BD also influence schizophrenia and major depression, albeit with variable effects⁶². SNPs in and around the MHC locus reached genome-wide significance for BD for the first time. However, unlike in schizophrenia, we found no influence of *C4* structural alleles or gene expression⁶³. Rather, the association was driven by variation outside the classical MHC locus, with the index SNP (*rs13195402*) being a missense

variant in *BTN2A1*, a brain-expressed gene⁶⁴ encoding a plasma membrane protein.

The genetic correlation of BD with other psychiatric disorders was consistent with previous reports^{65,66}. Our results also corroborate previous genetic and clinical evidence of associations between BD and sleep disturbances⁶⁷, problematic alcohol use⁶⁸ and smoking⁶⁹. While the genome-wide genetic correlations with these traits were modest ($r_g = -0.05$ to 0.35), MiXeR estimated that, for all traits, more than 55% of trait-influencing variants also influence BD (Fig. 3). Taken together, these results point to shared biology as one possible explanation for the high prevalence of substance use in BD. However, excluding genetic variants associated with both traits, MR analyses suggested that smoking is also a putatively 'causal' risk factor for BD, while BD has no effect on smoking, consistent with a previous report⁷⁰. (We use the word 'causal' with caution here as we consider MR an exploratory analysis to identify potentially modifiable risk factors that warrant more detailed investigations to understand their complex relationship with BD.) In contrast, MR indicated that BD had bidirectional 'causal' relationships with problematic alcohol use, longer sleep duration and mood instability. Insights into the relationship of such behavioral correlates with BD may have future impact on clinical decision-making in the prophylaxis or management of the disorder. Higher educational attainment has previously been associated with BD in epidemiological studies^{55,56}, while lower educational attainment has been associated with schizophrenia and major depression^{71,72}. Here, educational attainment had a significant positive effect on risk of BD and vice versa. Interestingly, MiXeR estimated that almost all variants that influence BD also influence educational attainment. The substantial genetic overlap observed between BD and the other phenotypes suggests that many variants likely influence multiple phenotypes, which may be differentiated by phenotype-specific effect size distributions among the shared influencing variants.

The integration of eQTL data with our GWAS results yielded 15 high-confidence genes for which there was converging evidence that their association with BD is mediated via gene expression. Among these were *HTR6*, encoding a serotonin receptor targeted by antipsychotics and antidepressants⁷³, and *MCHRI* (melanin-concentrating hormone receptor 1), encoding a target of the antipsychotic haloperidol⁷³. We note that, for both of these genes, their top eQTLs have opposite directions of effect on gene expression in the brain and blood, possibly playing a role in the tissue-specific gene regulation influencing BD⁷⁴. BD was associated with decreased expression of *FURIN*, a gene with a neurodevelopmental role that has already been the subject of functional genomics experiments in neuronal cells following its association with schizophrenia in GWAS⁷⁵. The top association in our GWAS was in the *TRANK1* locus on chromosome 3, which has previously been implicated in BD^{12,18,59}. Although BD-associated SNPs in this locus are known to regulate *TRANK1* expression⁷⁶, our eQTL analyses support a stronger but correlated regulation of *DCLK3*, located 87 kilobases (kb) upstream of *TRANK1* (refs. 43,77). Both *FURIN* and *DCLK3* also encode druggable proteins (although they are not targets for any current psychiatric medications)^{73,78}. These eQTL results provide promising BD candidate genes for functional follow-up experiments²⁹. While several of these are in genome-wide significant loci, many are not the closest gene to the index SNP, highlighting the value of probing underlying molecular mechanisms to prioritize the most likely causal genes in the loci.

GWAS signals were enriched in the gene targets of existing BD pharmacological agents, such as antipsychotics, mood stabilizers and antiepileptics. However, enrichment was also found in the targets of calcium channel blockers used to treat hypertension and GABA-receptor-targeting anesthetics (Supplementary Table 8). Calcium channel antagonists have long been investigated for the treatment of BD, without becoming an established therapeutic approach, and there is evidence that some antiepileptics have

calcium-channel-inhibiting effects^{79,80}. These results underscore the opportunity for repurposing some classes of drugs, particularly calcium channel antagonists, as potential BD treatments⁸¹.

BD associations were enriched in gene sets involving neuronal parts and synaptic signaling. Neuronal and synaptic pathways have been described in cross-disorder GWAS of multiple psychiatric disorders including BD^{82–84}. Dysregulation of such pathways has also been suggested by previous functional and animal studies⁸⁵. Analysis of single-cell gene expression data revealed enrichment in genes with high specificity of gene expression in neurons (both excitatory and inhibitory) of many brain regions, in particular the cortex and hippocampus. These findings are similar to those reported in GWAS data of schizophrenia⁸⁶ and major depressive disorder³⁸.

PRSs for BD explained on average 4.57% of phenotypic variance (liability scale) across European cohorts, although this varied in different waves of the BD GWAS, ranging from 6.6% in the PGC1 cohorts to 2.9% in the external biobank studies (Supplementary Fig. 7 and Supplementary Table 12). These results are in line with the h^2_{SNP} of BD per wave, which ranged from 24.6% (s.e. = 0.01) in PGC1 to 11.9% (s.e. = 0.01) in external studies (Supplementary Table 3). Some variability in h^2_{SNP} estimates may arise from the inclusion of cases from population biobanks, who may have more heterogeneous clinical presentations or less severe illness than patients with BD ascertained via inpatient or outpatient psychiatric clinics. Across the waves of clinically ascertained samples within the PGC, h^2_{SNP} and the R^2 of PRSs also varied, likely reflecting clinical and genetic heterogeneity in the type of BD cases ascertained; the PGC1 cohorts consisted mostly of BD I cases⁹, known to be the most heritable of the BD subtypes^{11,24}, while later waves included more individuals with BD II²⁴. Overall, the h^2_{SNP} of BD calculated from the meta-analysis summary statistics was 18% on the liability scale, a decrease of ~2% compared with the PGC2 GWAS²⁴, which may be due to the addition of cohorts with lower h^2_{SNP} estimates and heterogeneity between cohorts (Supplementary Table 3). However, despite differences in h^2_{SNP} and R^2 of PRSs per wave, the genetic correlation of BD between all waves was high (weighted mean $r_g = 0.94$, s.e. = 0.03), supporting our rationale for combining cases with different BD subtypes or ascertainment to increase power for discovery of risk variants. In Europeans, individuals in the top 10% of PRSs had an OR of 3.5 for BD, compared with individuals with average PRSs (middle decile), which translates into a modest absolute lifetime risk of the disorder (7% based on PRSs alone). While PRSs are invaluable tools in research settings, the current BD PRSs lack sufficient power to separate individuals into clinically meaningful risk categories, and therefore have no clinical utility at present^{47,88}. PRSs from this European BD meta-analysis yield higher R^2 values in diverse ancestry samples than PRSs based on any currently available BD GWAS within the same ancestry⁵⁹. However, performance still greatly lags behind that in Europeans, with ~2% variance explained in East Asian samples and substantially less in admixed African American samples, likely due to differences in allele frequencies and LD structures, consistent with previous studies^{89,90}. There is a pressing need for more and larger studies in other ancestry groups to ensure that any future clinical utility is broadly applicable. Exploiting the differences in LD structure between diverse ancestry samples will also assist in the fine-mapping of risk loci for BD.

Our analyses confirmed that BD is a highly polygenic disorder, with an estimated 8,600 variants explaining 90% of its h^2_{SNP} . Hence, many more SNPs than those identified here are expected to account for the common variant architecture underlying BD. This GWAS marks an inflection point in risk variant discovery, and we expect that, from this point forward, the addition of more samples will lead to a dramatic increase in genetic findings. Nevertheless, fewer genome-wide significant loci have been identified in BD than in a schizophrenia GWASs of comparable sample size⁶⁰. This may be due to the clinical and genetic heterogeneity that exists in BD.

Our GWAS of subtypes BD I and BD II identified additional associated loci. Consistent with previous findings²⁴, our analysis showed that the two subtypes were highly but imperfectly genetically correlated ($r_g = 0.85$), and that BD I is more genetically correlated with schizophrenia, while BD II has stronger genetic correlation with major depression. The subtypes are sufficiently similar to justify joint analysis as BD, but are not identical in their genetic composition, and as such contribute to the genetic heterogeneity of BD⁹¹. We identified 13 loci passing genome-wide significance for BD I, and one for BD II, which did not reach significance in the main BD GWAS, further illustrating the partially differing genetic composition of the two subtypes. Understanding the shared and distinct genetic components of BD subtypes and symptoms requires detailed phenotyping efforts in large cohorts and is an important area for future psychiatric genetics research.

In summary, these new data advance our understanding of the biological etiology of BD and prioritize a set of candidate genes for functional follow-up experiments. Several lines of evidence converge on the involvement of calcium channel signaling, providing a promising avenue for future therapeutic development.

Online content

Any methods, additional references, Nature Research reporting summaries, source data, extended data, supplementary information, acknowledgements, peer review information; details of author contributions and competing interests; and statements of data and code availability are available at <https://doi.org/10.1038/s41588-021-00857-4>.

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Niamh Mullins^{1,2,235}✉, **Andreas J. Forstner**^{3,4,5,235}, **Kevin S. O'Connell**^{6,7}, **Brandon Coombes**⁸, **Jonathan R. I. Coleman**^{9,10}, **Zhen Qiao**¹¹, **Thomas D. Als**^{12,13,14}, **Tim B. Bigdeli**^{15,16}, **Sigrid Børte**^{17,18,19}, **Julien Bryois**²⁰, **Alexander W. Charney**², **Ole Kristian Drange**^{21,22}, **Michael J. Gandal**²³, **Saskia P. Hagenaars**^{9,10}, **Masashi Ikeda**²⁴, **Nolan Kamitaki**^{25,26}, **Minsoo Kim**²³, **Kristi Krebs**²⁷, **Georgia Panagiotaropoulou**²⁸, **Brian M. Schilder**^{1,29,30,31}, **Laura G. Sloofman**¹, **Stacy Steinberg**³², **Vassily Trubetsky**²⁸, **Bendik S. Winsvold**^{19,33}, **Hong-Hee Won**³⁴, **Liliya Abramova**³⁵, **Kristina Adorjan**^{36,37}, **Esben Agerbo**^{14,38,39}, **Mariam Al Eissa**⁴⁰, **Diego Albani**⁴¹, **Ney Alliey-Rodriguez**^{42,43},

Adebayo Anjorin⁴⁴, Veneri Antilla⁴⁵, Anastasia Antoniou⁴⁶, Swapnil Awasthi²⁸, Ji Hyun Baek⁴⁷, Marie Bækvad-Hansen^{14,48}, Nicholas Bass⁴⁰, Michael Bauer⁴⁹, Eva C. Beins³, Sarah E. Bergen²⁰, Armin Birner⁵⁰, Carsten Bøcker Pedersen^{14,38,39}, Erlend Bøen⁵¹, Marco P. Boks⁵², Rosa Bosch^{53,54,55,56}, Murielle Brum⁵⁷, Ben M. Brumpton¹⁹, Nathalie Brunkhorst-Kanaan⁵⁷, Monika Budde³⁶, Jonas Bybjerg-Grauholm^{14,48}, William Byerley⁵⁸, Murray Cairns⁵⁹, Miquel Casas^{53,54,55,56}, Pablo Cervantes⁶⁰, Toni-Kim Clarke⁶¹, Cristiana Cruceanu^{60,62}, Alfredo Cuellar-Barboza^{63,64}, Julie Cunningham⁶⁵, David Curtis^{66,67}, Piotr M. Czerski⁶⁸, Anders M. Dale⁶⁹, Nina Dalkner⁵⁰, Friederike S. David³, Franziska Degenhardt^{3,70}, Srdjan Djurovic^{71,72}, Amanda L. Dobbyn^{1,2}, Athanassios Douzenis⁴⁶, Torbjørn Elvsåshagen^{18,73,74}, Valentina Escott-Price⁷⁵, I. Nicol Ferrier⁷⁶, Alessia Fiorentino⁴⁰, Tatiana M. Foroud⁷⁷, Liz Forty⁷⁵, Josef Frank⁷⁸, Oleksandr Frei^{6,18}, Nelson B. Freimer^{23,79}, Louise Frisén⁸⁰, Katrin Gade^{36,81}, Julie Garnham⁸², Joel Gelernter^{83,84,85}, Marianne Giørtz Pedersen^{14,38,39}, Ian R. Gizer⁸⁶, Scott D. Gordon⁸⁷, Katherine Gordon-Smith⁸⁸, Tiffany A. Greenwood⁸⁹, Jakob Grove^{12,13,14,90}, José Guzman-Parra⁹¹, Kyooseob Ha⁹², Magnus Haraldsson⁹³, Martin Hautzinger⁹⁴, Urs Heilbronner³⁶, Dennis Hellgren²⁰, Stefan Herms^{3,95,96}, Per Hoffmann^{3,95,96}, Peter A. Holmans⁷⁵, Laura Huckins^{1,2}, Stéphane Jamain^{97,98}, Jessica S. Johnson^{1,2}, Janos L. Kalman^{36,37,99}, Yoichiro Kamatani^{100,101}, James L. Kennedy^{102,103,104,105}, Sarah Kittel-Schneider^{57,106}, James A. Knowles^{107,108}, Manolis Kogevinas¹⁰⁹, Maria Koromina¹¹⁰, Thorsten M. Kranz⁵⁷, Henry R. Kranzler^{111,112}, Michiaki Kubo¹¹³, Ralph Kupka^{114,115,116}, Steven A. Kushner¹¹⁷, Catharina Lavebratt^{118,119}, Jacob Lawrence¹²⁰, Markus Leber¹²¹, Heon-Jeong Lee¹²², Phil H. Lee¹²³, Shawn E. Levy¹²⁴, Catrin Lewis⁷⁵, Calwing Liao^{125,126}, Susanne Lucae⁶², Martin Lundberg^{118,119}, Donald J. MacIntyre¹²⁷, Sigurdur H. Magnusson³², Wolfgang Maier¹²⁸, Adam Maihofer⁸⁹, Dolores Malaspina^{1,2}, Eirini Maratou¹²⁹, Lina Martinsson⁸⁰, Manuel Mattheisen^{12,13,14,106,130}, Steven A. McCarroll^{25,26}, Nathaniel W. McGregor¹³¹, Peter McGuffin⁹, James D. McKay¹³², Helena Medeiros¹⁰⁸, Sarah E. Medland⁸⁷, Vincent Millischer^{118,119}, Grant W. Montgomery¹¹, Jennifer L. Moran^{25,133}, Derek W. Morris¹³⁴, Thomas W. Mühleisen^{4,95}, Niamh O'Brien⁴⁰, Claire O'Donovan⁸², Loes M. Olde Loohuis^{23,79}, Lilijana Oruc¹³⁵, Sergi Papiol^{36,37}, Antonio F. Pardiñas⁷⁵, Amy Perry⁸⁸, Andrea Pfennig⁴⁹, Evgenia Porichi⁴⁶, James B. Potash¹³⁶, Digby Quested^{137,138}, Towfique Raj^{1,29,30,31}, Mark H. Rapaport¹³⁹, J. Raymond DePaulo¹³⁶, Eline J. Regeer¹⁴⁰, John P. Rice¹⁴¹, Fabio Rivas⁹¹, Margarita Rivera^{142,143}, Julian Roth¹⁰⁶, Panos Roussos^{1,2,29}, Douglas M. Ruderfer¹⁴⁴, Cristina Sánchez-Mora^{53,54,56,145}, Eva C. Schulte^{36,37}, Fanny Senner^{36,37}, Sally Sharp⁴⁰, Paul D. Shilling⁸⁹, Engilbert Sigurdsson^{93,146}, Lea Sirignano⁷⁸, Claire Slaney⁸², Olav B. Smeland^{6,7}, Daniel J. Smith¹⁴⁷, Janet L. Sobell¹⁴⁸, Christine Søholm Hansen^{14,48}, Maria Soler Artigas^{53,54,56,145}, Anne T. Spijker¹⁴⁹, Dan J. Stein¹⁵⁰, John S. Strauss¹⁰², Beata Świątkowska¹⁵¹, Chikashi Terao¹⁰¹, Thorgeir E. Thorgeirsson³², Claudio Toma^{152,153,154}, Paul Tooney⁵⁹, Evangelia-Eirini Tsermpini¹¹⁰, Marquis P. Vawter¹⁵⁵, Helmut Vedder¹⁵⁶, James T. R. Walters⁷⁵, Stephanie H. Witt⁷⁸, Simon Xi¹⁵⁷, Wei Xu¹⁵⁸, Jessica Mei Kay Yang⁷⁵, Allan H. Young^{159,160}, Hannah Young¹, Peter P. Zandi¹³⁶, Hang Zhou^{83,84}, Lea Zillich⁷⁸, HUNT All-In Psychiatry*, Rolf Adolfsson¹⁶¹, Ingrid Agartz^{51,130,162}, Martin Alda^{82,163}, Lars Alfredsson¹⁶⁴, Gulja Babadjanova¹⁶⁵, Lena Backlund^{118,119}, Bernhard T. Baune^{166,167,168}, Frank Bellivier^{169,170}, Susanne Bengesser⁵⁰, Wade H. Berrettini¹⁷¹, Douglas H. R. Blackwood⁶¹, Michael Boehnke¹⁷², Anders D. Børglum^{14,173,174}, Gerome Breen^{9,10}, Vaughan J. Carr¹⁷⁵, Stanley Catts¹⁷⁶, Aiden Corvin¹⁷⁷, Nicholas Craddock⁷⁵, Udo Dannlowski¹⁶⁶, Dimitris Dikeos¹⁷⁸, Tõnu Esko^{26,27,179,180}, Bruno Etain^{169,170}, Panagiotis Ferentinos^{9,46}, Mark Frye⁶⁴, Janice M. Fullerton^{152,153}, Micha Gawlik¹⁰⁶, Elliot S. Gershon^{42,181},

Fernando S. Goes¹³⁶, Melissa J. Green^{152,175}, Maria Grigoriu-Serbanescu¹⁸², Joanna Hauser⁶⁸, Frans Henskens⁵⁹, Jan Hillert⁸⁰, Kyung Sue Hong⁴⁷, David M. Hougaard^{14,48}, Christina M. Hultman²⁰, Kristian Hveem^{19,183}, Nakao Iwata²⁴, Assen V. Jablensky¹⁸⁴, Ian Jones⁷⁵, Lisa A. Jones⁸⁸, René S. Kahn^{2,52}, John R. Kelsoe⁸⁹, George Kirov⁷⁵, Mikael Landén^{20,185}, Marion Leboyer^{97,98,186}, Cathryn M. Lewis^{9,10,187}, Qingqin S. Li¹⁸⁸, Jolanta Lissowska¹⁸⁹, Christine Lochner¹⁹⁰, Carmel Loughland⁵⁹, Nicholas G. Martin^{87,191}, Carol A. Mathews¹⁹², Fermin Mayoral⁹¹, Susan L. McElroy¹⁹³, Andrew M. McIntosh^{127,194}, Francis J. McMahon¹⁹⁵, Ingrid Melle^{6,196}, Patricia Michie⁵⁹, Lili Milani²⁷, Philip B. Mitchell¹⁷⁵, Gunnar Morken^{21,197}, Ole Mors^{14,198}, Preben Bo Mortensen^{12,14,38,39}, Bryan Mowry¹⁷⁶, Bertram Müller-Myhsok^{62,199,200}, Richard M. Myers¹²⁴, Benjamin M. Neale^{25,45,179}, Caroline M. Nievergelt^{89,201}, Merete Nordentoft^{14,202}, Markus M. Nöthen³, Michael C. O'Donovan⁷⁵, Ketil J. Oedegaard^{203,204}, Tomas Olsson²⁰⁵, Michael J. Owen⁷⁵, Sara A. Paciga²⁰⁶, Chris Pantelis²⁰⁷, Carlos Pato¹⁰⁸, Michele T. Pato¹⁰⁸, George P. Patrinos^{110,208,209}, Roy H. Perlis^{210,211}, Danielle Posthuma^{212,213}, Josep Antoni Ramos-Quiroga^{53,54,55,56}, Andreas Reif⁵⁷, Eva Z. Reininghaus⁵⁰, Marta Ribasés^{53,54,56,145}, Marcella Rietschel⁷⁸, Stephan Ripke^{25,28,45}, Guy A. Rouleau^{126,214}, Takeo Saito²⁴, Ulrich Schall⁵⁹, Martin Schalling^{118,119}, Peter R. Schofield^{152,153}, Thomas G. Schulze^{36,78,81,136,215}, Laura J. Scott¹⁷², Rodney J. Scott⁵⁹, Alessandro Serretti²¹⁶, Cynthia Shannon Weickert^{152,175,217}, Jordan W. Smoller^{25,133,218}, Hreinn Stefansson³², Kari Stefansson^{32,219}, Eystein Stordal^{220,221}, Fabian Streit⁷⁸, Patrick F. Sullivan^{20,222,223}, Gustavo Turecki²²⁴, Arne E. Vaaler²²⁵, Eduard Vieta²²⁶, John B. Vincent¹⁰², Irwin D. Waldman²²⁷, Thomas W. Weickert^{152,175,217}, Thomas Werge^{14,228,229,230}, Naomi R. Wray^{11,231}, John-Anker Zwart^{18,19,33}, Joanna M. Biernacka^{8,64}, John I. Nurnberger²³², Sven Cichon^{3,4,95,96}, Howard J. Edenberg^{77,233}, Eli A. Stahl^{1,2,179,236}, Andrew McQuillin^{40,236}, Arianna Di Florio^{75,223,236}, Roel A. Ophoff^{23,79,117,234,236} and Ole A. Andreassen^{6,7,236}  

¹Department of Genetics and Genomic Sciences, Icahn School of Medicine at Mount Sinai, New York, NY, USA. ²Department of Psychiatry, Icahn School of Medicine at Mount Sinai, New York, NY, USA. ³Institute of Human Genetics, University of Bonn, School of Medicine and University Hospital Bonn, Bonn, Germany. ⁴Institute of Neuroscience and Medicine (INM-1), Research Centre Jülich, Jülich, Germany. ⁵Centre for Human Genetics, University of Marburg, Marburg, Germany. ⁶Division of Mental Health and Addiction, Oslo University Hospital, Oslo, Norway. ⁷NORMENT, University of Oslo, Oslo, Norway. ⁸Department of Health Sciences Research, Mayo Clinic, Rochester, MN, USA. ⁹Social, Genetic and Developmental Psychiatry Centre, King's College London, London, UK. ¹⁰NIHR Maudsley BRC, King's College London, London, UK. ¹¹Institute for Molecular Bioscience, The University of Queensland, Brisbane, Queensland, Australia. ¹²iSEQ, Center for Integrative Sequencing, Aarhus University, Aarhus, Denmark. ¹³Department of Biomedicine - Human Genetics, Aarhus University, Aarhus, Denmark. ¹⁴iPSYCH, The Lundbeck Foundation Initiative for Integrative Psychiatric Research, Aarhus, Denmark. ¹⁵Department of Psychiatry and Behavioral Sciences, SUNY Downstate Health Sciences University, Brooklyn, NY, USA. ¹⁶VA NY Harbor Healthcare System, Brooklyn, NY, USA. ¹⁷Research and Communication Unit for Musculoskeletal Health, Division of Clinical Neuroscience, Oslo University Hospital, Oslo, Norway. ¹⁸Institute of Clinical Medicine, University of Oslo, Oslo, Norway. ¹⁹K. G. Jebsen Center for Genetic Epidemiology, Department of Public Health and Nursing, Faculty of Medicine and Health Sciences, Norwegian University of Science and Technology, Trondheim, Norway. ²⁰Department of Medical Epidemiology and Biostatistics, Karolinska Institutet, Stockholm, Sweden. ²¹Department of Mental Health, Faculty of Medicine and Health Sciences, Norwegian University of Science and Technology (NTNU), Trondheim, Norway. ²²Department of Østmarka, Division of Mental Health Care, St Olavs Hospital, Trondheim University Hospital, Trondheim, Norway. ²³Department of Psychiatry and Biobehavioral Science, Semel Institute, David Geffen School of Medicine, University of California, Los Angeles, Los Angeles, CA, USA. ²⁴Department of Psychiatry, School of Medicine, Fujita Health University, Toyoake, Japan. ²⁵Stanley Center for Psychiatric Research, Broad Institute, Cambridge, MA, USA. ²⁶Department of Genetics, Harvard Medical School, Boston, MA, USA. ²⁷Estonian Genome Center, Institute of Genomics, University of Tartu, Tartu, Estonia. ²⁸Department of Psychiatry and Psychotherapy, Charité - Universitätsmedizin, Berlin, Germany. ²⁹Department of Neuroscience, Icahn School of Medicine at Mount Sinai, New York, NY, USA. ³⁰Ronald M. Loeb Center for Alzheimer's Disease, Icahn School of Medicine at Mount Sinai, New York, NY, USA. ³¹Estelle and Daniel Maggin Department of Neurology, Icahn School of Medicine at Mount Sinai, New York, NY, USA. ³²deCODE Genetics/Amgen, Reykjavik, Iceland. ³³Department of Research, Innovation and Education, Division of Clinical Neuroscience, Oslo University Hospital, Oslo, Norway. ³⁴Samsung Advanced Institute for Health Sciences and Technology (SAIHST), Samsung Medical Center, Sungkyunkwan University, Seoul, South Korea. ³⁵Russian Academy of Medical Sciences, Mental Health Research Center, Moscow, Russian Federation. ³⁶Institute of Psychiatric Phenomics and Genomics (IPPG), University Hospital, LMU Munich, Munich, Germany. ³⁷Department of Psychiatry and Psychotherapy, University Hospital, LMU Munich, Munich, Germany. ³⁸National Centre for Register-Based Research, Aarhus University, Aarhus, Denmark. ³⁹Centre for Integrated Register-Based Research, Aarhus University, Aarhus, Denmark. ⁴⁰Division of Psychiatry, University College London, London, UK. ⁴¹Department of Neuroscience, Istituto Di Ricerche Farmacologiche Mario Negri IRCCS, Milan, Italy. ⁴²Department of Psychiatry and Behavioral Neuroscience, University of Chicago, Chicago, IL, USA. ⁴³Northwestern University, Chicago, IL, USA. ⁴⁴Psychiatry, Berkshire Healthcare NHS Foundation Trust, Bracknell, UK. ⁴⁵Analytic and Translational Genetics Unit, Massachusetts General Hospital, Boston, MA, USA. ⁴⁶2nd Department of Psychiatry, Attikon General Hospital, National and Kapodistrian University of Athens, Athens, Greece. ⁴⁷Department of Psychiatry, Samsung Medical Center, School of Medicine, Sungkyunkwan University, Seoul, South Korea. ⁴⁸Center for Neonatal Screening, Department for Congenital

Disorders, Statens Serum Institut, Copenhagen, Denmark. ⁴⁹Department of Psychiatry and Psychotherapy, University Hospital Carl Gustav Carus, Technische Universität Dresden, Dresden, Germany. ⁵⁰Department of Psychiatry and Psychotherapeutic Medicine, Medical University of Graz, Graz, Austria. ⁵¹Department of Psychiatric Research, Diakonhjemmet Hospital, Oslo, Norway. ⁵²Psychiatry, Brain Center UMC Utrecht, Utrecht, the Netherlands. ⁵³Instituto de Salud Carlos III, Biomedical Network Research Centre on Mental Health (CIBERSAM), Madrid, Spain. ⁵⁴Department of Psychiatry, Hospital Universitari Vall d'Hebron, Barcelona, Spain. ⁵⁵Department of Psychiatry and Forensic Medicine, Universitat Autònoma de Barcelona, Barcelona, Spain. ⁵⁶Psychiatric Genetics Unit, Group of Psychiatry Mental Health and Addictions, Vall d'Hebron Research Institut (VHIR), Universitat Autònoma de Barcelona, Barcelona, Spain. ⁵⁷Department of Psychiatry, Psychosomatic Medicine and Psychotherapy, University Hospital Frankfurt, Frankfurt am Main, Germany. ⁵⁸Psychiatry, University of California San Francisco, San Francisco, CA, USA. ⁵⁹University of Newcastle, Newcastle, New South Wales, Australia. ⁶⁰Mood Disorders Program, Department of Psychiatry, McGill University Health Center, Montreal, Quebec, Canada. ⁶¹Division of Psychiatry, University of Edinburgh, Edinburgh, UK. ⁶²Department of Translational Research in Psychiatry, Max Planck Institute of Psychiatry, Munich, Germany. ⁶³Department of Psychiatry, Universidad Autonoma de Nuevo Leon, Monterrey, Mexico. ⁶⁴Department of Psychiatry and Psychology, Mayo Clinic, Rochester, MN, USA. ⁶⁵Department of Laboratory Medicine and Pathology, Mayo Clinic, Rochester, MN, USA. ⁶⁶Centre for Psychiatry, Queen Mary University of London, London, UK. ⁶⁷UCL Genetics Institute, University College London, London, UK. ⁶⁸Department of Psychiatry, Laboratory of Psychiatric Genetics, Poznan University of Medical Sciences, Poznan, Poland. ⁶⁹Center for Multimodal Imaging and Genetics, Departments of Neurosciences, Radiology, and Psychiatry, University of California, San Diego, CA, USA. ⁷⁰Department of Child and Adolescent Psychiatry, Psychosomatics and Psychotherapy, University Hospital Essen, University of Duisburg-Essen, Duisburg, Germany. ⁷¹Department of Medical Genetics, Oslo University Hospital, Oslo, Norway. ⁷²NORMENT, Department of Clinical Science, University of Bergen, Bergen, Norway. ⁷³Department of Neurology, Oslo University Hospital, Oslo, Norway. ⁷⁴NORMENT, KG Jebsen Centre for Psychosis Research, Oslo University Hospital, Oslo, Norway. ⁷⁵Medical Research Council Centre for Neuropsychiatric Genetics and Genomics, Division of Psychological Medicine and Clinical Neurosciences, Cardiff University, Cardiff, UK. ⁷⁶Academic Psychiatry, Newcastle University, Newcastle upon Tyne, UK. ⁷⁷Department of Medical and Molecular Genetics, Indiana University, Indianapolis, IN, USA. ⁷⁸Department of Genetic Epidemiology in Psychiatry, Central Institute of Mental Health, Medical Faculty Mannheim, Heidelberg University, Mannheim, Germany. ⁷⁹Center for Neurobehavioral Genetics, Semel Institute for Neuroscience and Human Behavior, Los Angeles, CA, USA. ⁸⁰Department of Clinical Neuroscience, Karolinska Institutet, Stockholm, Sweden. ⁸¹Department of Psychiatry and Psychotherapy, University Medical Center Göttingen, Göttingen, Germany. ⁸²Department of Psychiatry, Dalhousie University, Halifax, Nova Scotia, Canada. ⁸³Department of Psychiatry, Yale School of Medicine, New Haven, CT, USA. ⁸⁴Veterans Affairs Connecticut Healthcare System, West Haven, CT, USA. ⁸⁵Departments of Genetics and Neuroscience, Yale University School of Medicine, New Haven, CT, USA. ⁸⁶Department of Psychological Sciences, University of Missouri, Columbia, MO, USA. ⁸⁷Genetics and Computational Biology, QIMR Berghofer Medical Research Institute, Brisbane, Queensland, Australia. ⁸⁸Psychological Medicine, University of Worcester, Worcester, UK. ⁸⁹Department of Psychiatry, University of California San Diego, La Jolla, CA, USA. ⁹⁰Bioinformatics Research Centre, Aarhus University, Aarhus, Denmark. ⁹¹Mental Health Department, University Regional Hospital, Biomedicine Institute (IBIMA), Málaga, Spain. ⁹²Department of Psychiatry, Seoul National University College of Medicine, Seoul, South Korea. ⁹³Landspítali University Hospital, Reykjavík, Iceland. ⁹⁴Department of Psychology, Eberhard Karls Universität Tübingen, Tübingen, Germany. ⁹⁵Department of Biomedicine, University of Basel, Basel, Switzerland. ⁹⁶Institute of Medical Genetics and Pathology, University Hospital Basel, Basel, Switzerland. ⁹⁷Neuropsychiatrie Translationnelle, Inserm U955, Créteil, France. ⁹⁸Faculté de Santé, Université Paris Est, Créteil, France. ⁹⁹International Max Planck Research School for Translational Psychiatry (IMPRS-TP), Munich, Germany. ¹⁰⁰Laboratory of Complex Trait Genomics, Department of Computational Biology and Medical Sciences, Graduate School of Frontier Sciences, The University of Tokyo, Tokyo, Japan. ¹⁰¹Laboratory for Statistical and Translational Genetics, RIKEN Center for Integrative Medical Sciences, Yokohama, Japan. ¹⁰²Campbell Family Mental Health Research Institute, Centre for Addiction and Mental Health, Toronto, Ontario, Canada. ¹⁰³Neurogenetics Section, Centre for Addiction and Mental Health, Toronto, Ontario, Canada. ¹⁰⁴Department of Psychiatry, University of Toronto, Toronto, Ontario, Canada. ¹⁰⁵Institute of Medical Sciences, University of Toronto, Toronto, Ontario, Canada. ¹⁰⁶Department of Psychiatry, Psychosomatics and Psychotherapy, Center of Mental Health, University Hospital Würzburg, Würzburg, Germany. ¹⁰⁷Cell Biology, SUNY Downstate Medical Center College of Medicine, Brooklyn, NY, USA. ¹⁰⁸Institute for Genomic Health, SUNY Downstate Medical Center College of Medicine, Brooklyn, NY, USA. ¹⁰⁹ISGlobal, Barcelona, Spain. ¹¹⁰Laboratory of Pharmacogenomics and Individualized Therapy, Department of Pharmacy, School of Health Sciences, University of Patras, Patras, Greece. ¹¹¹Mental Illness Research, Education and Clinical Center, Crescenz VAMC, Philadelphia, PA, USA. ¹¹²Center for Studies of Addiction, University of Pennsylvania Perelman School of Medicine, Philadelphia, PA, USA. ¹¹³RIKEN Center for Integrative Medical Sciences, Yokohama, Japan. ¹¹⁴Psychiatry, Altrecht, Utrecht, the Netherlands. ¹¹⁵Psychiatry, GGZ inGeest, Amsterdam, the Netherlands. ¹¹⁶Psychiatry, VU Medisch Centrum, Amsterdam, the Netherlands. ¹¹⁷Department of Psychiatry, Erasmus MC, University Medical Center Rotterdam, Rotterdam, the Netherlands. ¹¹⁸Department of Molecular Medicine and Surgery, Karolinska Institutet, Stockholm, Sweden. ¹¹⁹Center for Molecular Medicine, Karolinska University Hospital, Stockholm, Sweden. ¹²⁰Psychiatry, North East London NHS Foundation Trust, Ilford, UK. ¹²¹Clinic for Psychiatry and Psychotherapy, University Hospital Cologne, Cologne, Germany. ¹²²Department of Psychiatry, Korea University College of Medicine, Seoul, South Korea. ¹²³Psychiatric and Neurodevelopmental Genetics Unit, Center for Genomic Medicine, Massachusetts General Hospital and Harvard Medical School, Boston, MA, USA. ¹²⁴HudsonAlpha Institute for Biotechnology, Huntsville, AL, USA. ¹²⁵Department of Human Genetics, McGill University, Montréal, Quebec, Canada. ¹²⁶Montreal Neurological Institute and Hospital, McGill University, Montréal, Quebec, Canada. ¹²⁷Division of Psychiatry, Centre for Clinical Brain Sciences, The University of Edinburgh, Edinburgh, UK. ¹²⁸Department of Psychiatry and Psychotherapy, University of Bonn, Bonn, Germany. ¹²⁹Clinical Biochemistry Laboratory, Attikon General Hospital, Medical School, National and Kapodistrian University of Athens, Athens, Greece. ¹³⁰Department of Clinical Neuroscience, Centre for Psychiatry Research, Karolinska Institutet, Stockholm, Sweden. ¹³¹Systems Genetics Working Group, Department of Genetics, Stellenbosch University, Stellenbosch, South Africa. ¹³²Genetic Cancer Susceptibility Group, International Agency for Research on Cancer, Lyon, France. ¹³³Department of Psychiatry, Massachusetts General Hospital, Boston, MA, USA. ¹³⁴Centre for Neuroimaging and Cognitive Genomics (NICOG), National University of Ireland Galway, Galway, Ireland. ¹³⁵Medical Faculty, School of Science and Technology, University Sarajevo, Sarajevo, Bosnia and Herzegovina. ¹³⁶Department of Psychiatry and Behavioral Sciences, Johns Hopkins University School of Medicine, Baltimore, MD, USA. ¹³⁷Oxford Health NHS Foundation Trust, Warneford Hospital, Oxford, UK. ¹³⁸Department of Psychiatry, University of Oxford, Warneford Hospital, Oxford, UK. ¹³⁹Department of Psychiatry and Behavioral Sciences, Emory University School of Medicine, Atlanta, GA, USA. ¹⁴⁰Outpatient Clinic for Bipolar Disorder, Altrecht, Utrecht, the Netherlands. ¹⁴¹Department of Psychiatry, Washington University in Saint Louis, Saint Louis, MO, USA. ¹⁴²Department of Biochemistry and Molecular Biology II, Faculty of Pharmacy, University of Granada, Granada, Spain. ¹⁴³Institute of Neurosciences, Biomedical Research Center (CIBM), University of Granada, Granada, Spain. ¹⁴⁴Medicine, Psychiatry, Biomedical Informatics, Vanderbilt University Medical Center, Nashville, TN, USA. ¹⁴⁵Department of Genetics, Microbiology and Statistics, Faculty of Biology, Universitat de Barcelona, Barcelona, Spain. ¹⁴⁶Faculty of Medicine, Department of Psychiatry, School of Health Sciences, University of Iceland, Reykjavik, Iceland. ¹⁴⁷Institute of Health and Wellbeing, University of Glasgow, Glasgow, UK. ¹⁴⁸Psychiatry and the Behavioral Sciences, University of Southern California, Los Angeles, CA, USA. ¹⁴⁹Mood Disorders, PsyQ, Rotterdam, the Netherlands. ¹⁵⁰SAMRC Unit on Risk and Resilience in Mental Disorders, Department of Psychiatry and Neuroscience Institute, University of Cape Town, Cape Town, South Africa. ¹⁵¹Department of Environmental Epidemiology, Nofer Institute of Occupational Medicine, Lodz, Poland. ¹⁵²Neuroscience Research Australia, Sydney, New South Wales, Australia. ¹⁵³School of Medical Sciences, University of New South Wales, Sydney, New South Wales, Australia. ¹⁵⁴Centro de Biología Molecular Severo Ochoa, Universidad Autónoma de Madrid and CSIC, Madrid, Spain.

¹⁵⁵Department of Psychiatry and Human Behavior, School of Medicine, University of California, Irvine, Irvine, CA, USA. ¹⁵⁶Psychiatry, Psychiatrisches Zentrum Nordbaden, Wiesloch, Germany. ¹⁵⁷Computational Sciences Center of Emphasis, Pfizer Global Research and Development, Cambridge, MA, USA. ¹⁵⁸Dalla Lana School of Public Health, University of Toronto, Toronto, Ontario, Canada. ¹⁵⁹Department of Psychological Medicine, Institute of Psychiatry, Psychology and Neuroscience, King's College London, London, UK. ¹⁶⁰South London and Maudsley NHS Foundation Trust, Bethlem Royal Hospital, Beckenham, UK. ¹⁶¹Department of Clinical Sciences, Psychiatry, Umeå University Medical Faculty, Umeå, Sweden. ¹⁶²NORMENT, KG Jebsen Centre for Psychosis Research, Division of Mental Health and Addiction, Institute of Clinical Medicine and Diakonhjemmet Hospital, University of Oslo, Oslo, Norway. ¹⁶³National Institute of Mental Health, Klecany, Czech Republic. ¹⁶⁴Institute of Environmental Medicine, Karolinska Institutet, Stockholm, Sweden. ¹⁶⁵Institute of Pulmonology, Russian State Medical University, Moscow, Russian Federation. ¹⁶⁶Department of Psychiatry, University of Münster, Münster, Germany. ¹⁶⁷Department of Psychiatry, Melbourne Medical School, The University of Melbourne, Melbourne, Victoria, Australia. ¹⁶⁸The Florey Institute of Neuroscience and Mental Health, The University of Melbourne, Parkville, Victoria, Australia. ¹⁶⁹Université de Paris, INSERM, Optimisation Thérapeutique en Neuropsychopharmacologie, UMRS 1144, Paris, France. ¹⁷⁰APHP Nord, DMU Neurosciences, Département de Psychiatrie et de Médecine Addictologique, GHU Saint Louis-Lariboisière-Fernand Widal, Paris, France. ¹⁷¹Psychiatry, University of Pennsylvania, Philadelphia, PA, USA. ¹⁷²Center for Statistical Genetics and Department of Biostatistics, University of Michigan, Ann Arbor, MI, USA. ¹⁷³Department of Biomedicine and the iSEQ Center, Aarhus University, Aarhus, Denmark. ¹⁷⁴Center for Genomics and Personalized Medicine, CGPM, Aarhus, Denmark. ¹⁷⁵School of Psychiatry, University of New South Wales, Sydney, New South Wales, Australia. ¹⁷⁶University of Queensland, Brisbane, Queensland, Australia. ¹⁷⁷Neuropsychiatric Genetics Research Group, Department of Psychiatry and Trinity Translational Medicine Institute, Trinity College Dublin, Dublin, Ireland. ¹⁷⁸1st Department of Psychiatry, Eginition Hospital, National and Kapodistrian University of Athens, Athens, Greece. ¹⁷⁹Medical and Population Genetics, Broad Institute, Cambridge, MA, USA. ¹⁸⁰Division of Endocrinology, Children's Hospital Boston, Boston, MA, USA. ¹⁸¹Department of Human Genetics, University of Chicago, Chicago, IL, USA. ¹⁸²Biometric Psychiatric Genetics Research Unit, Alexandru Obregia Clinical Psychiatric Hospital, Bucharest, Romania. ¹⁸³HUNT Research Center, Department of Public Health and Nursing, Faculty of Medicine and Health Sciences, Norwegian University of Science and Technology, Trondheim, Norway. ¹⁸⁴University of Western Australia, Nedlands, Western Australia, Australia. ¹⁸⁵Institute of Neuroscience and Physiology, University of Gothenburg, Gothenburg, Sweden. ¹⁸⁶Department of Psychiatry and Addiction Medicine, Assistance Publique - Hôpitaux de Paris, Paris, France. ¹⁸⁷Department of Medical and Molecular Genetics, King's College London, London, UK. ¹⁸⁸Neuroscience Therapeutic Area, Janssen Research and Development, LLC, Titusville, NJ, USA. ¹⁸⁹Cancer Epidemiology and Prevention, M. Skłodowska-Curie National Research Institute of Oncology, Warsaw, Poland. ¹⁹⁰SA MRC Unit on Risk and Resilience in Mental Disorders, Department of Psychiatry, Stellenbosch University, Stellenbosch, South Africa. ¹⁹¹School of Psychology, The University of Queensland, Brisbane, Queensland, Australia. ¹⁹²Department of Psychiatry and Genetics Institute, University of Florida, Gainesville, FL, USA. ¹⁹³Research Institute, Lindner Center of HOPE, Mason, OH, USA. ¹⁹⁴Centre for Cognitive Ageing and Cognitive Epidemiology, University of Edinburgh, Edinburgh, UK. ¹⁹⁵Human Genetics Branch, Intramural Research Program, National Institute of Mental Health, Bethesda, MD, USA. ¹⁹⁶Division of Mental Health and Addiction, University of Oslo, Institute of Clinical Medicine, Oslo, Norway. ¹⁹⁷Psychiatry, St Olavs University Hospital, Trondheim, Norway. ¹⁹⁸Psychosis Research Unit, Aarhus University Hospital - Psychiatry, Risskov, Denmark. ¹⁹⁹Munich Cluster for Systems Neurology (SyNergy), Munich, Germany. ²⁰⁰University of Liverpool, Liverpool, UK. ²⁰¹Research/Psychiatry, Veterans Affairs San Diego Healthcare System, San Diego, CA, USA. ²⁰²Mental Health Services in the Capital Region of Denmark, Mental Health Center Copenhagen, University of Copenhagen, Copenhagen, Denmark. ²⁰³Division of Psychiatry, Haukeland Universitetssjukehus, Bergen, Norway. ²⁰⁴Faculty of Medicine and Dentistry, University of Bergen, Bergen, Norway. ²⁰⁵Department of Clinical Neuroscience and Center for Molecular Medicine, Karolinska Institutet at Karolinska University Hospital, Solna, Sweden. ²⁰⁶Human Genetics and Computational Biomedicine, Pfizer Global Research and Development, Groton, CT, USA. ²⁰⁷University of Melbourne, Melbourne, Victoria, Australia. ²⁰⁸Department of Pathology, College of Medicine and Health Sciences, United Arab Emirates University, Al-Ain, United Arab Emirates. ²⁰⁹Zayed Center of Health Sciences, United Arab Emirates University, Al-Ain, United Arab Emirates. ²¹⁰Psychiatry, Harvard Medical School, Boston, MA, USA. ²¹¹Division of Clinical Research, Massachusetts General Hospital, Boston, MA, USA. ²¹²Department of Complex Trait Genetics, Center for Neurogenetics and Cognitive Research, Amsterdam Neuroscience, Vrije Universiteit Amsterdam, Amsterdam, the Netherlands. ²¹³Department of Clinical Genetics, Amsterdam Neuroscience, Vrije Universiteit Medical Center, Amsterdam, the Netherlands. ²¹⁴Department of Neurology and Neurosurgery, Faculty of Medicine, McGill University, Montreal, Quebec, Canada. ²¹⁵Department of Psychiatry and Behavioral Sciences, SUNY Upstate Medical University, Syracuse, NY, USA. ²¹⁶Department of Biomedical and NeuroMotor Sciences, University of Bologna, Bologna, Italy. ²¹⁷Department of Neuroscience, SUNY Upstate Medical University, Syracuse, NY, USA. ²¹⁸Psychiatric and Neurodevelopmental Genetics Unit (PNGU), Massachusetts General Hospital, Boston, MA, USA. ²¹⁹Faculty of Medicine, University of Iceland, Reykjavik, Iceland. ²²⁰Department of Psychiatry, Hospital Namsos, Namsos, Norway. ²²¹Department of Neuroscience, Norges Teknisk Naturvitenskapelige Universitet Fakultet for Naturvitenskap og Teknologi, Trondheim, Norway. ²²²Department of Genetics, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA. ²²³Department of Psychiatry, University of North Carolina at Chapel Hill, Chapel Hill, NC, USA. ²²⁴Department of Psychiatry, McGill University, Montreal, Quebec, Canada. ²²⁵Department of Psychiatry, Sankt Olavs Hospital Universitetssykehuset i Trondheim, Trondheim, Norway. ²²⁶Clinical Institute of Neuroscience, Hospital Clinic, University of Barcelona, IDIBAPS, CIBERSAM, Barcelona, Spain. ²²⁷Department of Psychology, Emory University, Atlanta, GA, USA. ²²⁸Institute of Biological Psychiatry, Mental Health Services, Copenhagen University Hospital, Copenhagen, Denmark. ²²⁹Department of Clinical Medicine, University of Copenhagen, Copenhagen, Denmark. ²³⁰Center for GeoGenetics, GLOBE Institute, University of Copenhagen, Copenhagen, Denmark. ²³¹Queensland Brain Institute, The University of Queensland, Brisbane, Queensland, Australia. ²³²Psychiatry, Indiana University School of Medicine, Indianapolis, IN, USA. ²³³Biochemistry and Molecular Biology, Indiana University School of Medicine, Indianapolis, IN, USA. ²³⁴Department of Human Genetics, David Geffen School of Medicine, University of California Los Angeles, Los Angeles, CA, USA. ²³⁵These authors contributed equally: Niamh Mullins, Andreas J. Forstner. ²³⁶These authors jointly supervised this work: Eli A. Stahl, Andrew McQuillin, Arianna Di Florio, Roel A. Ophoff, Ole A. Andreassen. *A list of authors and their affiliations appears at the end of the paper.

[✉]e-mail: niamh.mullins@mssm.edu; ole.andreassen@medisin.uio.no

HUNT All-In Psychiatry

Bendik S. Winsvold^{19,33}, Eystein Stordal^{220,221}, Gunnar Morken^{21,197}, John-Anker Zwart^{18,19,33}, Ole Kristian Drange^{21,22} and Sigrid Børte^{17,18,19}

A list of members and their affiliations appears in the Supplementary Information.

Methods

Sample description. The meta-analysis sample comprises 57 cohorts collected in Europe, North America and Australia, totaling 41,917 BD cases and 371,549 controls of European descent (Supplementary Table 1). The total effective n , equivalent to an equal number of cases and controls in each cohort ($4 \times n_{\text{cases}} \times n_{\text{controls}} / (n_{\text{cases}} + n_{\text{controls}})$), is 101,962. For 52 cohorts, individual-level genotype and phenotype data were shared with the PGC. Cohorts have been added to the PGC in five waves (PGC1 (ref. ⁷), PGC2 (ref. ²⁴), PGC PsychChip, PGC3 and External studies); all cohorts from previous PGC BD GWASs were included. The source and inclusion/exclusion criteria for cases and controls for each cohort are described in the Supplementary Note. Cases were required to meet international consensus criteria (DSM-IV, ICD-9 or ICD-10) for a lifetime diagnosis of BD, established using structured diagnostic instruments from assessments by trained interviewers, clinician-administered checklists or medical record review. In most cohorts, controls were screened for the absence of lifetime psychiatric disorders and randomly selected from the population. For five cohorts (iPSYCH³⁰, deCODE genetics³¹, Estonian Biobank³², HUNT³³ and UK Biobank³⁴), GWAS summary statistics for BD were shared with the PGC. In these cohorts, BD cases were ascertained using ICD codes or self-report during a nurse interview, and the majority of controls were screened for the absence of psychiatric disorders via ICD codes. Follow-up analyses included four non-European BD case-control cohorts, two from East Asia (Japan³⁹ and Korea⁹²) and two admixed African American cohorts^{22,93}, providing a total of 5,847 cases and 65,588 controls. These BD cases were ascertained using international consensus criteria (DSM-IV)^{22,93} through psychiatric interviews (Supplementary Note).

Genotyping, quality control and imputation. For 52 cohorts internal to the PGC, genotyping was performed following local protocols and genotypes were called using standard genotype calling software from commercial sources (Affymetrix and Illumina). Subsequently, standardized quality control, imputation and statistical analyses were performed centrally using Rapid Imputation for Consortias Pipeline (RICOPILI; version 2018_Nov_23.001)⁹⁴, separately for each cohort. Briefly, the quality control parameters for retaining SNPs and subjects were: SNP missingness < 0.05 (before sample removal), subject missingness < 0.02, autosomal heterozygosity deviation ($F_{\text{het}} < 0.2$), SNP missingness < 0.02 (after sample removal), difference in SNP missingness between cases and controls < 0.02, SNP Hardy-Weinberg equilibrium ($P > 10 \times 10^{-10}$ in psychiatric cases and $P > 10 \times 10^{-6}$ in controls). Relatedness was calculated across cohorts using identity by descent and one of each pair of related individuals ($\pi_{\text{hat}} > 0.2$) was excluded. Principal components (PCs) were generated using genotyped SNPs in each cohort separately using EIGENSTRAT v6.1.4 (ref. ⁹⁵). On the basis of visual inspection of plots of PCs for each dataset (which were all of European descent according to self-report/clinical data), we excluded samples to obtain more clearly homogeneous datasets. Genotype imputation was performed using the prephasing/imputation stepwise approach implemented in Eagle v2.3.5 (ref. ⁹⁶) and Minimac3 (ref. ⁹⁷) to the Haplotype Reference Consortium (HRC) reference panel v1.0 (ref. ⁹⁸). Data on the X chromosome were available for 50 cohorts internal to the PGC and one external cohort (HUNT), and the X chromosome was imputed to the HRC reference panel in males and females separately within each cohort. The five external cohorts were processed by the collaborating research teams using comparable procedures and imputed to the HRC or a custom reference panel as appropriate. Full details of the genotyping, quality control and imputation for each of these cohorts are available in the Supplementary Note. Identical individuals between PGC cohorts and the Estonian Biobank and UK Biobank cohorts were detected using genotype-based checks (https://personal.broadinstitute.org/sripke/share_links/zpXkV8INxUg9bayDpLToG4g58TMtjN_PGC_SCZ_w3.01718.76) and removed from PGC cohorts.

GWAS. For PGC cohorts, GWASs were conducted within each cohort using an additive logistic regression model in PLINK v1.90 (ref. ⁹⁹), covarying for PCs 1–5 and any others as required. Association analyses of the X chromosome were conducted in males and females separately using the same procedures, with males coded as 0 or 2 for 0 or 1 copies of the reference allele. Results from males and females were then meta-analyzed within each cohort. For external cohorts, GWASs were conducted by the collaborating research teams using comparable procedures (Supplementary Note). To control test statistic inflation at SNPs with low minor allele frequency (MAF) in small cohorts, SNPs were retained only if cohort MAF was > 1% and minor allele count was > 10 in either cases or controls (whichever had smaller n). There was no evidence of stratification artifacts or uncontrolled inflation of test statistics in the results from any cohort ($\lambda_{\text{GC}} = 0.97$ –1.05; Supplementary Table 1). Meta-analysis of GWAS summary statistics was conducted using an inverse-variance-weighted fixed-effects model in METAL (version 2011-03-25)¹⁰⁰ across 57 cohorts for the autosomes (41,917 BD cases and 371,549 controls) and 51 cohorts for the X chromosome (35,691 BD cases and 96,731 controls). A genome-wide significant locus was defined as the region around a SNP with $P < 5 \times 10^{-8}$, with LD $r^2 > 0.1$, within a 3,000-kb window. Regional association plots and forest plots of the index SNPs for all genome-wide significant loci are presented in Supplementary Data 1 and 2, respectively.

Overlap of loci with other psychiatric disorders. Genome-wide significant loci for BD were assessed for overlap with genome-wide significant loci for other psychiatric disorders, using the largest available GWAS results for major depression⁶¹, schizophrenia⁶⁰, attention deficit/hyperactivity disorder¹⁰¹, post-traumatic stress disorder¹⁰², lifetime anxiety disorder¹⁰³, Tourette's syndrome¹⁰⁴, anorexia nervosa¹⁰⁵, alcohol use disorder or problematic alcohol use⁶⁸, autism spectrum disorder¹⁰⁶, mood disorders⁵¹ and the cross-disorder GWAS of the PGC⁶⁶. The boundaries of the genome-wide significant loci were calculated in the original publications. Overlap of loci was calculated using bedtools v2.29.2 (ref. ¹⁰⁷).

Enrichment analyses. P values quantifying the degree of association of genes and gene sets with BD were calculated using MAGMA v1.08 (ref. ³⁷), implemented in FUMA v1.3.6a (refs. ^{64,108}). Gene-based tests were performed for 19,576 genes (Bonferroni-corrected P -value threshold = 2.55×10^{-6}). A total of 11,858 curated gene sets including at least 10 genes from MSigDB V7.0 were tested for association with BD (Bonferroni-corrected P -value threshold = 4.22×10^{-6}). Competitive gene-set tests were conducted correcting for gene size, variant density and LD within and between genes. Tissue-set enrichment analyses were also performed using MAGMA implemented in FUMA, to test for enrichment of association signal in genes expressed in 54 tissue types from GTEx V8 (Bonferroni-corrected P -value threshold = 9.26×10^{-4})^{64,108}.

For single-cell enrichment analyses, publicly available single-cell RNA-sequencing data were compiled from five studies of the adult human and mouse brain^{86,109–112}. The mean expression for each gene in each cell type was computed from the single-cell expression data (if not provided). For the Zeisel dataset¹⁰⁹, we used the mean expression at level 4 (39 cell types from 19 regions for the mouse nervous system). For the Saunders dataset¹¹⁰, we computed the mean expression of the different classes in each of the 9 different brain regions sampled (88 cell types in total). We filtered out any genes with nonunique names, genes not expressed in any cell types, non-protein-coding genes and, for mouse datasets, genes that had no expert-curated 1:1 orthologs between mouse and human (Mouse Genome Informatics, The Jackson Laboratory, version 11/22/2016, <http://www.informatics.jax.org/downloads/reports/index.html#homology>), resulting in 16,472 genes. Gene expression was then scaled to a total of 1 million unique molecular identifiers (or transcripts per million) for each cell type/tissue. Using a previously described method³⁸, a metric of gene expression specificity was calculated by dividing the expression of each gene in each cell type by the total expression of that gene in all cell types, leading to values ranging from 0 to 1 for each gene (0 meaning that the gene is not expressed in that cell type and 1 meaning that all of the expression of the gene is in that cell type). We then selected the top 10% most specific genes for each cell type/tissue for enrichment analysis. MAGMA v1.08 (ref. ³⁷) was used to test gene-set enrichment using GWAS summary statistics, covarying for gene size, gene density, mean sample size for tested SNPs per gene, the inverse of the minor allele counts per gene and the log of these metrics. We excluded any SNPs with INFO score < 0.6, with MAF < 1% or with estimated odds ratio > 25 or smaller than 1/25, as well as SNPs located in the MHC region (chr6:25–34 Mb). We set a window of 35 kb upstream to 10 kb downstream of the gene coordinates to compute gene-level association statistics and used the European reference panel from phase 3 of the 1000 Genomes Project as the reference population¹¹³. We then used MAGMA to test whether the 10% most specific genes (with an expression of at least 1 transcript per million or 1 unique molecular identifier per million) for each cell type/tissue were associated with BD. The P -value threshold for significance was $P < 9.1 \times 10^{-3}$, representing a 5% false discovery rate across datasets.

Further gene-set analyses were performed restricted to genes targeted by drugs, assessing individual drugs and grouping drugs with similar actions. This approach has been described previously⁴¹. Gene-level and gene-set analyses were performed in MAGMA v1.08 (ref. ³⁷). Gene boundaries were defined using build 37 reference data from the National Center for Biotechnology Information, available on the MAGMA website (<https://ctg.cncr.nl/software/magma>), extended 35 kb upstream and 10 kb downstream to include regulatory regions outside the transcribed region. Gene-level association statistics were defined as the aggregate of the mean and the lowest variant-level P value within the gene boundary, converted to a Z value. Gene sets were defined comprising the targets of each drug in the Drug-Gene Interaction database DGIdb v2 (ref. ³⁹) and in the Psychoactive Drug Screening Database Ki DB⁴⁰, both downloaded in June 2016⁴¹. Analyses were performed using competitive gene-set analyses in MAGMA. Results from the drug-set analysis were then grouped according to the Anatomical Therapeutic Chemical class of the drug⁴¹. Only drug classes with at least ten valid drug gene sets within them were analyzed. Drug-class analysis was performed using enrichment curves. All drug gene sets were ranked by their association in the drug-set analysis, and then for a given drug class an enrichment curve was drawn scoring a 'hit' if the drug gene set was within the class, or a 'miss' if it was outside the class. The area under the curve was calculated, and a P value for this was calculated using the Wilcoxon Mann-Whitney test comparing drug gene sets within the class to drug gene sets outside the class⁴¹. Multiple testing was controlled using a Bonferroni-corrected significance threshold of $P < 5.60 \times 10^{-5}$ for drug-set analysis and $P < 7.93 \times 10^{-4}$ for drug-class analysis, accounting for 893 drug sets and 63 drug classes tested.

eQTL integrative analysis. A TWAS was conducted using the precomputed gene expression weights from PsychENCODE data (1,321 brain samples)⁴³, available online with the FUSION software⁴². For genes with significant *cis*-SNP heritability (13,435 genes), FUSION software (vOct 1, 2019) was used to test whether SNPs influencing gene expression are also associated with BD (Bonferroni-corrected P -value threshold $< 3.72 \times 10^{-6}$). For regions including a TWAS-significant gene, TWAS fine-mapping of the region was conducted using FOCUS (fine-mapping of causal gene sets, v0.6.10)⁴⁴. Regions were defined using the correlation matrix of predicted effects on gene expression around TWAS-significant genes⁴⁴. A PIP was assigned to each gene for being causal for the observed TWAS association signal. Based on the PIP of each gene and a null model, whereby no gene in the region is causal for the TWAS signal, the 90%-credible gene set for each region was computed⁴⁴.

SMR (v1.03)^{45,46} was applied to further investigate putative causal relationships between SNPs and BD via gene expression. SMR was performed using eQTL summary statistics from the eQTLGen (31,684 blood samples)⁴⁷ and PsychENCODE⁴³ consortia. SMR analysis is limited to transcripts with at least one significant *cis*-eQTL ($P < 5 \times 10^{-8}$) in each dataset (15,610 in eQTLGen; 10,871 in PsychENCODE). The Bonferroni-corrected significance threshold was $P < 3.20 \times 10^{-6}$ and $P < 4.60 \times 10^{-6}$ for eQTLGen and PsychENCODE, respectively. The significance threshold for the HEIDI test was $P_{\text{HEIDI}} \geq 0.01$ (ref.⁴⁶). While the results of TWAS and SMR indicate an association between BD and gene expression, a nonsignificant HEIDI test additionally indicates either a direct causal role or a pleiotropic effect of the BD-associated SNPs on gene expression.

C4 imputation. To investigate the MHC (chr6:24–34 Mb on hg19), the alleles of *C4* genes (*C4A* and *C4B*) were imputed in 47 PGC cohorts for which individual-level genotype data were accessible, totaling 32,749 BD cases and 53,370 controls. The imputation reference panel comprised 2,530 reference haplotypes of MHC SNPs and *C4* alleles, generated using a sample of 1,265 individuals with whole-genome sequence data, from the Genomic Psychiatry cohort¹¹⁴. Briefly, imputation of *C4* as a multiallelic variant was performed using Beagle v4.1 (refs.^{115,116}), using SNPs from the MHC region that were also in the haplotype reference panel. Within the Beagle pipeline, the reference panel was first converted to bref format. We used the conform-gt tool to perform strand-flipping and filtering of specific SNPs for which the strand remained ambiguous. Beagle was run using default parameters with two key exceptions: we used the GRCh37 PLINK recombination map, and we set the output to include genotype probability (that is, GP field in VCF) for correct downstream probabilistic estimation of *C4A* and *C4B* joint dosages. The output consisted of dosage estimates for each of the common *C4* structural haplotypes for each individual. The five most common structural forms of the *C4A/C4B* locus (BS, AL, AL-BS, AL-BL and AL-AL) could be inferred with reasonably high accuracy (generally $0.70 < r^2 < 1.00$). The imputed *C4* alleles were tested for association with BD in a joint logistic regression that included terms for dosages of the five most common *C4* structural haplotypes (AL-BS, AL-BL, AL-AL, BS and AL), rs13195402 genotype (top lead SNP in the MHC) and PCs as per the GWAS. The genetically regulated expression of *C4A* was predicted from the imputed *C4* alleles using a model previously described⁶³. Predicted *C4A* expression was tested for association with BD in a joint logistic regression that included predicted *C4A* expression, rs13195402 genotype (top lead SNP in the MHC) and PCs as per the GWAS.

Polygenic risk scoring. PRSs from our GWAS meta-analysis were tested for association with BD in individual cohorts, using a discovery GWAS where the target cohort was left out of the meta-analysis. Briefly, the GWAS results from each discovery GWAS were pruned for LD using the P -value-informed clumping method in PLINK v1.90 (ref.⁹⁹; $r^2 \leq 0.1$ within a 500-kb window) based on the LD structure of the HRC reference panel⁹⁸. Subsets of SNPs were selected from the results below nine increasingly liberal P -value thresholds (GWAS P_i ; 5×10^{-8} , 1×10^{-4} , 1×10^{-3} , 0.01, 0.05, 0.1, 0.2, 0.5, 1). Sets of alleles, weighted by their log odds ratios from the discovery GWAS, were summed into PRSs for each individual in the target datasets, using PLINK v1.90 implemented via RICOPILI^{94,99}. PRSs were tested for association with BD in the target dataset using logistic regression, covarying for PCs as per the GWAS in each cohort. PRSs were tested in the external cohorts by the collaborating research teams using comparable procedures. The variance explained by the PRSs (R^2) was converted to the liability scale to account for the proportion of cases in each target dataset, using a BD population prevalence of 2% and 1%¹¹⁷. The weighted average R^2 values were calculated using the effective n for each cohort. The odds ratios for BD for individuals in the top decile of PRSs compared with those in the lowest decile and middle decile were calculated in the 52 datasets internal to the PGC. To assess cross-ancestry performance, PRSs generated from the meta-analysis results were tested for association with BD using similar methods in a Japanese sample⁹⁹, a Korean sample⁹² and two admixed African American samples. Full details of the QC, imputation and analysis of these samples are in the Supplementary Note.

LDSC. LDSC³⁵ was used to estimate the h_{SNP}^2 of BD from GWAS summary statistics. h_{SNP}^2 was converted to the liability scale, using a lifetime BD prevalence of 2% and 1%. LDSC bivariate genetic correlations attributable to genome-wide SNPs

(r_g) were estimated with 255 human diseases and traits from published GWASs and 514 GWASs of phenotypes in the UK Biobank from LD Hub⁴⁸. Adjusting for the number of traits tested, the Bonferroni-corrected P -value thresholds were $P < 1.96 \times 10^{-4}$ and $P < 9.73 \times 10^{-5}$, respectively.

MiXeR. We applied causal mixture models^{49,50} to the GWAS summary statistics, using MiXeR v1.3. MiXeR provides univariate estimates of the proportion of non-null SNPs ('polygenicity') and the variance of effect sizes of non-null SNPs ('discoverability') in each phenotype. For each SNP, i , univariate MiXeR models its additive genetic effect of allele substitution, β_i , as a point-normal mixture, $\beta_i = (1 - \pi_1) N(0, 0) + \pi_1 N(0, \sigma_{\beta_i}^2)$, where π_1 represents the proportion of non-null SNPs ('polygenicity') and $\sigma_{\beta_i}^2$ represents variance of effect sizes of non-null SNPs ('discoverability'). Then, for each SNP, j , MiXeR incorporates LD information and allele frequencies for $M = 9,997,231$ SNPs extracted from 1000 Genomes phase 3 data to estimate the expected probability distribution of the signed test statistic, $z_j = \delta_j + \epsilon_j = N \sum_i \sqrt{H_i} r_{ij} \beta_i + \epsilon_j$, where N is sample size, H_i indicates heterozygosity of the i th SNP, r_{ij} indicates allelic correlation between the i th and j th SNPs and $\epsilon_j \sim N(0, \sigma_{\epsilon_j}^2)$ is the residual variance. Further, the three parameters, π_1 , $\sigma_{\beta_i}^2$ and $\sigma_{\epsilon_j}^2$, are fitted by direct maximization of the likelihood function. The optimization is based on a set of approximately 600,000 SNPs, obtained by selecting a random set of 2,000,000 SNPs with MAF of 5% or higher, followed by LD pruning at LD $r^2 = 0.8$ threshold. The random SNP selection and full optimization procedure are repeated 20 times to obtain the means and standard errors of model parameters. The log-likelihood figures show individual curves for each of the 20 runs, each shifted vertically so that the best log-likelihood point is shown at the zero ordinate.

The total number of trait-influencing variants is estimated as $M\pi_1$, where $M = 9,997,231$ gives the number of SNPs in the reference panel. MiXeR Venn diagrams report the effective number of influencing variants, $\eta M\pi_1$, where η is a fixed number, $\eta = 0.319$, which gives the fraction of influencing variants contributing to 90% of the trait's heritability (with rationale for this adjustment being that the remaining 68.1% of influencing variants are small and cumulatively explain only 10% of the trait's heritability). Phenotypic variance explained on average by an influencing genetic variant is calculated as $\frac{H \sigma_{\beta_i}^2}{M}$, where $\frac{H}{M} = \frac{1}{M} \sum_i H_i = 0.2075$ is the average heterozygosity across SNPs in the reference panel. Under the assumptions of the MiXeR model, SNP heritability is then calculated as $h_{\text{SNP}}^2 = M\pi_1 \times H \sigma_{\beta_i}^2$.

In the cross-trait analysis, MiXeR models additive genetic effects as a mixture of four components, representing null SNPs in both traits (π_0); SNPs with a specific effect on the first and on the second trait (π_1 and π_2 , respectively); and SNPs with nonzero effect on both traits (π_{12}). In the last component, MiXeR

models the variance-covariance matrix as $\Sigma_{12} = \begin{bmatrix} \sigma_1^2 & \rho_{12} \sigma_1 \sigma_2 \\ \rho_{12} \sigma_1 \sigma_2 & \sigma_2^2 \end{bmatrix}$, where

ρ_{12} indicates correlation of effect sizes within the shared component, and σ_1^2 and σ_2^2 correspond to the discoverability parameter estimated in the univariate analysis of the two traits. These components are then plotted in Venn diagrams. After fitting parameters of the model, the Dice coefficient of polygenic overlap is then calculated as $\frac{2\pi_{12}}{\pi_1 + 2\pi_{12} + \pi_2}$, and genetic correlation is calculated as

$r_g = \frac{\rho_{12} \pi_{12}}{\sqrt{(\pi_1 + \pi_{12})(\pi_2 + \pi_{12})}}$. The fraction of influencing variants with concordant

effect direction is calculated as twice the multivariate normal cumulative distribution function at point (0, 0) for the bivariate normal distribution with zero mean and variance-covariance matrix Σ_{12} . All code is available online (<https://github.com/precimed/mixer>).

MR. We selected 17 traits associated with BD in clinical or epidemiological studies for MR to dissect their relationship with BD (Supplementary Note). Bidirectional GSMR⁵¹ analyses were performed between BD and the traits of interest using GWAS summary statistics, implemented in GCTA software (v1.93.1f beta). The instrumental variables were selected by a clumping procedure internal to the GSMR software with parameters: --gwas-thresh 5×10^{-8} --clump-r2 0.01. Traits with fewer than 10 instrumental variables available were excluded from the GSMR analyses to avoid conducting underpowered tests⁵¹, resulting in 10 traits tested (Bonferroni-corrected P -value threshold $< 2.5 \times 10^{-3}$). The HEIDI-outlier test was applied to test for horizontal pleiotropy ($P_{\text{HEIDI}} < 0.01$)⁵¹. For comparison, the MR analyses were also performed using the inverse-variance-weighted regression method, implemented via the TwoSampleMR R package, using the instrumental variables selected by GSMR^{118,119}. To further investigate horizontal pleiotropy, the MR Egger intercept test was conducted using the TwoSampleMR package^{118,119} and MR-PRESSO software was used to perform the global test and the distortion test¹²⁰.

BD subtypes. GWAS meta-analyses were conducted for BD I (25,060 cases, 449,978 controls from 55 cohorts, effective $n = 64,802$) and BD II (6,781 cases, 364,075 controls from 31 cohorts, effective $n = 22,560$; Supplementary Table 1) using the same procedures described for the main GWAS. BD subtypes were defined based on international consensus criteria (DSM-IV, ICD-9 or ICD-10), established using structured diagnostic instruments from assessments by

trained interviewers, clinician-administered checklists or medical record review. In the external biobank cohorts, BD subtypes were defined using ICD codes (Supplementary Note). LDSC³⁵ was used to estimate the h_{SNP}^2 of each subtype, and the genetic correlation between the subtypes. The difference between the LDSC h_{SNP}^2 estimates for BD I and BD II was tested for deviation from 0 using the block jackknife²¹. The LDSC genetic correlation (r_g) was tested for difference from 1 by calculating a chi-square statistic corresponding to the estimated r_g as $(r_g - 1)/s.e.^2$.

Reporting Summary. Further information on research design is available in the Nature Research Reporting Summary linked to this article.

Data availability

GWAS summary statistics are publicly available on the PGC website (<https://www.med.unc.edu/pgc/results-and-downloads>). Individual-level data are accessible through collaborative analysis proposals to the Bipolar Disorder Working Group of the PGC (<https://www.med.unc.edu/pgc/shared-methods/how-to/>). This study included some publicly available datasets accessed through dbGaP (PGC bundle [phs001254](https://www.ncbi.nlm.nih.gov/bioproject/1254)) and the HRC reference panel v1.0 (<http://www.haplotype-reference-consortium.org/home>). Databases used: Drug–Gene Interaction Database DGIdb v2 (<https://www.dgidb.org>); Psychoactive Drug Screening Database Ki DB (<https://pdsp.unc.edu/databases/kidb.php>); DrugBank 5.0 (<https://www.drugbank.ca>); LD Hub (<http://ldsc.broadinstitute.org>); FUMA (<https://fuma.ctglab.nl>).

Code availability

All software used is publicly available at the URLs or references cited.

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Author contributions

Writing group: N.M., A.J.F., K.S.O.C., B.C., J.R.I.C., J.M.B., J.I.N., S. Cichon, H.J.E., E.A.S., A. McQuillin, A.D.F., R.A.O., O.A.A. PGC BD PI group: A.J.F., M.I., H.-H.W., D.C., R.A., I.A., M.A., L. Alfredsson, G. Babadjanova, L.B., B.T.B., F.B., S. Bengesser, W.H.B., D.H.R.B., M. Boehnke, A.D.B., G. Breen, V.J.C., S. Catts, A.C., N.C., U.D., D.D., T. Esko, B.E., P.F., M.F., J.M.F., M.G., E.S.G., F.S.G., M. J. Green, M.G.-S., J. Hauser, F.H., J. Hillert, K.S.H., D.M.H., C.M.H., K. Hveem, N.I., A.V.J., I.J., L.A.J., R.S.K., J.R.K., G.K., M. Landén, M. Leboyer, C.M.L., Q.S.L., J. Lissowska, C. Lochner, C. Loughland, N.G.M., C.A.M., F.M., S.L.M., A.M.M., F.J.M., I.M., P. Michie, L. Milani, P. B. Mitchell, G.M., O.M., P. B. Mortensen, B.M., B.M.-M., R.M.M., B.M.N., C.M.N., M.N., M.M.N., M.C.O'D., K.J.O., T.O., M.J.O., S.A.P., C. Pantelis, C. Pato, M.T.P., G.P.P., R.H.P., D.P., J.A.R.-Q., A.R., E.Z.R., M. Ribasés, M. Rietschel, S.R., G.A.R., T.S., U.S., M.S., P.R.S., T.G.S., L.J.S., R.J.S., A.S., C.S.W., J.W.S., H.S., K.S., E. Stordal, F. Streit, P.F.S., G.T., A.E.V., E.V., J.B.V., I.D.W., T.W.W., T.W., N.R.W., J.-A.Z., J.M.B., J.I.N., S. Cichon, H.J.E., E.A.S., A. McQuillin, A.D.F., R.A.O., O.A.A. Bioinformatics: N.M., A.J.F., J.R.I.C., S. Børte, M.J. Gandal, M. Kim, B.M.S., L.G.S., B.S.W., H.-H.W., N.A.-R., S.E.B., B.M.B., V.E.-P., S.H., P.A.H., Y.K., M. Koromina, M. Kubo, M. Leber, P.H.L., C. Liao, L.M.O.L., T.R., P.R., P.D.S., M.S.A., C. Terao, T.E.T., S.X., H.Y., P.P.Z., S. Bengesser, G. Breen, P.F., E.S.G., Q.S.L., G.A.R., H.S., T.W., E.A.S. Clinical: O.K.D., M.I., L. Abramova, K.A., E.A., N.A.-R., A. Anjorin, A. Antoniou, J.H.B., N.B., M. Bauer, A.B., C.B.P., E.B., M.P.B., R.B., M. Brum, N.B.-K., M. Budde, W.B., M. Cairns, M. Casas, P.C., A.C.-B., D.C., P.M.C., N.D., A.D., T. Elvsåshagen, L. Forty, L. Frisén, K.G., J. Garnham, M.G.P., I.R.G., K.G.-S., J. Grove, J.G.-P., K. Ha, M. Haraldsson, M. Hautzinger, U.H., D.H., J. L. Kalman, J. L. Kennedy, S.K.-S., M. Kogevinas, T.M.K., R.K., S.A.K., J. Lawrence, H.-J.L., C. Lewis, S.L., M. Lundberg, D.J.M., W.M., D.M., L. Martinsson, M.M., P. McGuffin, H.M., V.M., C.O'D., L.O., S.P., A. Perry, A. Pfennig, E.P., J.B.P., D.Q., M.H.R., J.R.D., E.J.R., J.P.R., F.R., J.R., E.C.S., F. Sennar, E. Sigurdsson, L.S., C.S., O.B.S., D. J. Smith, J.L.S., A.T.S., J.S.S., B.S., P.T., M.P.V., H.V., A.H.Y., L.Z., HUNT All-In Psychiatry, R.A., I.A., M.A., G. Babadjanova, L.B., B.T.B., F.B., S. Bengesser, D.H.R.B., A.D.B., A.C., N.C., U.D., D.B., B.E., P.F., M.F., M.G., E.S.G., F.S.G., M. J. Green, M.G.-S., J. Hauser, K.S.H., N.I., I.J., L.A.J., R.S.K., G.K., M. Landén, C.M.L., J. Lissowska, N.G.M., C.A.M., F.M., S.L.M., A.M.M., I.M., P. B. Mitchell, G.M., O.M., P. B. Mortensen, M.C.O'D., K.J.O., M.J.O., C. Pato, M.T.P., R.H.P., J.A.R.-Q., A.R., E.Z.R., M. Rietschel, T.S., T.G.S., A.S., C.S.W., J.W.S., E. Stordal, F. Streit, A.E.V., E.V., J.B.V., I.D.W., T.W.W., T.W., N.R.W., J.-A.Z., M. Quillin, A.D.F. Genomic assays/data generation: A.J.F., M.I., E.A., M.A.E., D.A., M.B.-H., E.C.B., C.B.P., J.B.-G., M. Cairns, T.-K.C., C.C., J.C., F.S.D., F.D., S.D., A.F., J.E., N.B.F., J. Gelernter, M.G.P., P.H., S.J., Y.K., H.R.K., M. Kubo, S.E.L., C. Liao, E.M., N.W.M., J.D.M., G.W.M., J.L.M., D.W.M., T.W.M., N.O'B., M. Rivera, C.S.-M., S. Sharp, C.S.H., C. Terao, C. Toma, E.-E.T., S.H.W., HUNT All-In Psychiatry, G. Breen, A.C., T. Esko, J.M.F., E.S.G., D.M.H., N.I., F.J.M., L. Milani,

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Competing interests

T.E.T., S. Steinberg, H.S. and K.S. are employed by deCODE Genetics/Amgen. Multiple additional authors work for pharmaceutical or biotechnology companies in a manner directly analogous to academic coauthors and collaborators. A.H.Y. has given paid lectures and served on advisory boards relating to drugs used in affective and related disorders for several companies (AstraZeneca, Eli Lilly, Lundbeck, Sunovion, Servier, Livanova, Janssen, Allergan, Bionomics and Sumitomo Dainippon Pharma), was Lead Investigator for Embolden Study (AstraZeneca), BCI Neuroplasticity study and Aripiprazole Mania Study, and is an investigator for Janssen, Lundbeck, Livanova and Compass. J.I.N. is an investigator for Janssen. P.E.S. reports the following potentially competing financial interests: Lundbeck (advisory committee), Pfizer (Scientific Advisory Board member) and Roche (grant recipient, speaker reimbursement). G. Breen reports consultancy and speaker fees from Eli Lilly and Illumina and grant funding from Eli Lilly. M. Landén has received speaker fees from Lundbeck. O.A.A. has received speaker fees from Lundbeck and Sunovion, and is a consultant to HealthLytix. J.A.R.-Q. was on the speakers bureau and/or acted as consultant for Eli Lilly, Janssen-Cilag, Novartis, Shire, Lundbeck, Almirall, Braingaze, Sincrolab and Rubió in the last 5 years. He also received travel awards (air tickets and hotel) for taking part in psychiatric meetings from Janssen-Cilag, Rubió, Shire and Eli Lilly. The Department of Psychiatry

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Additional information

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Correspondence and requests for materials should be addressed to N.M. or O.A.A.

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Software and code

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Data collection Ascertainment of all cohorts is described in the Supplementary Note. Data were generated at many sites using standard genotyping pipelines and calling softwares from commercial sources (Affymetrix and Illumina).

Data analysis Analyses were performed using the publicly available RICOPILI software suite (version 2018_Nov_23.001) which provides wrappers for standard genetic analysis software including PLINK v1.09, Minimac3, Eagle v2.3.5, EIGENSTRAT v6.1.4 and METAL (version 2011-03-25). Post-GWAS analyses were conducted using FUMA v1.3.6a, MAGMA v1.08, FUSION (vOct 1, 2019), FOCUS 0.6.10, SMR v1.03, LDSC v1.0.0, MiXeR v1.3, GSMR implemented in GCTA software v1.93.1f beta, bedtools v2.29.2, TwoSampleMR v0.5.4 and MR-PRESSO v1.0 R packages. All software is publicly available. Any additional analysis code is detailed in Methods and the Supplementary Note.

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Sample size	Sample size was not predetermined but all available cohorts of bipolar disorder cases and controls were included to maximize sample size. This study provides a 1.5-fold increase in effective sample size compared with previous GWAS of bipolar disorder, increasing statistical power to detect novel genetic associations.
Data exclusions	Predetermined phenotypic data exclusions, for both cases and controls, are detailed in the Supplementary Note. Genotype data exclusions were also predetermined and were performed for quality control; these included high missing call rate, high or low heterozygosity, inconsistent genotype versus clinical data sex, and ancestry outlier status based on visual inspection of genotype principal component analysis results. Phenotypic and genotypic exclusions were applied to the GWAS and all post-GWAS analyses.
Replication	All available cohorts of bipolar disorder cases and controls were included in the primary analysis and therefore we do not perform replication of genetic associations in independent cohorts here. Post-GWAS analyses were replicated using different statistical genetics methods or by integrating different publicly available biological datasets where possible.
Randomization	Samples/participants were allocated into experimental groups by clinical cohort (which included country of origin) and genotype data collection batches. Association analyses were performed in each dataset and meta-analyzed across datasets. Ancestry covariates derived from genotype principal components analysis were included in association tests, which were logistic regression.
Blinding	Case and control groups were assigned by trained researchers or interviewers during cohort ascertainment, or using predefined International Classification of Diseases codes applied to biobanks. Therefore case/ control status was fixed at ascertainment. Experimenters were not blind to case/ control status while performing statistical analyses of the genetic data as this was not practical.

Reporting for specific materials, systems and methods

We require information from authors about some types of materials, experimental systems and methods used in many studies. Here, indicate whether each material, system or method listed is relevant to your study. If you are not sure if a list item applies to your research, read the appropriate section before selecting a response.

Materials & experimental systems

n/a	Involved in the study
<input checked="" type="checkbox"/>	<input type="checkbox"/> Antibodies
<input checked="" type="checkbox"/>	<input type="checkbox"/> Eukaryotic cell lines
<input checked="" type="checkbox"/>	<input type="checkbox"/> Palaeontology and archaeology
<input checked="" type="checkbox"/>	<input type="checkbox"/> Animals and other organisms
<input type="checkbox"/>	<input checked="" type="checkbox"/> Human research participants
<input checked="" type="checkbox"/>	<input type="checkbox"/> Clinical data
<input checked="" type="checkbox"/>	<input type="checkbox"/> Dual use research of concern

Methods

n/a	Involved in the study
<input checked="" type="checkbox"/>	<input type="checkbox"/> ChIP-seq
<input checked="" type="checkbox"/>	<input type="checkbox"/> Flow cytometry
<input checked="" type="checkbox"/>	<input type="checkbox"/> MRI-based neuroimaging

Human research participants

Policy information about [studies involving human research participants](#)

Population characteristics	This is fully described in the Online Methods and associated Supplementary Note and Tables.
Recruitment	This is fully described in the Online Methods and associated Supplementary Note and Tables.
Ethics oversight	All local IRBs approved of this study. This is fully described in the Online Methods and associated Supplementary Note and Tables.

Note that full information on the approval of the study protocol must also be provided in the manuscript.