

Men's status and reproductive success in 33 nonindustrial societies: Effects of subsistence, marriage system, and reproductive strategy

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Edited by Kristen Hawkes, University of Utah, Salt Lake City, UT, and approved July 19, 2016 (received for review April 28, 2016)

Social status motivates much of human behavior. However, status may have been a relatively weak target of selection for much of human evolution if ancestral foragers tended to be more egalitarian. We test the “egalitarianism hypothesis” that status has a significantly smaller effect on reproductive success (RS) in foragers compared with nonforagers. We also test between alternative male reproductive strategies, in particular whether reproductive benefits of status are due to lower offspring mortality (parental investment) or increased fertility (mating effort). We performed a phylogenetic multilevel metaanalysis of 288 statistical associations between measures of male status (physical formidability, hunting ability, material wealth, political influence) and RS (mating success, wife quality, fertility, offspring mortality, and number of surviving offspring) from 46 studies in 33 nonindustrial societies. We found a significant overall effect of status on RS ($r = 0.19$), though this effect was significantly lower than for nonhuman primates ($r = 0.80$). There was substantial variation due to marriage system and measure of RS, in particular status associated with offspring mortality only in polygynous societies ($r = -0.08$), and with wife quality only in monogamous societies ($r = 0.15$). However, the effects of status on RS did not differ significantly by status measure or subsistence type: foraging, horticulture, pastoralism, and agriculture. These results suggest that traits that facilitate status acquisition were not subject to substantially greater selection with domestication of plants and animals, and are part of reproductive strategies that enhance fertility more than offspring well-being.

status | hierarchy | reproduction | egalitarianism | evolution

Social status is a fundamental human motive (1). Social status can be defined as relative access to contested resources within a group, particularly the deference or admiration of group members (2, 3). To gain status, individuals attempt to influence perceptions of their dominance (i.e., ability to inflict costs on others) or prestige (i.e., ability to confer benefits on others) (3, 4). Dominance and prestige can be difficult to disentangle because dominance can be a source of prestige and vice versa (5). The particular traits that underlie dominance and prestige may vary across groups, but dominance and prestige are often based on conspicuous, interindividual differences in body size (6); intelligence and skill (7); consumption (8); and generosity (9, 10). Relative to the status hierarchies of other primates, prestige contributes heavily to human status, because of our interdependence in production and reproduction (11, 12), and thus the value of social partners and mates who are strong, skilled, wealthy, or generous (3, 13). Furthermore, humans frequently cooperate against dominant individuals who act coercively (14, 15).

Traits that facilitate status acquisition, including a desire for status, can have positive fitness consequences. Group members relinquish or offer reproductive opportunities to high-status individuals, to avoid costs from competition or gain reciprocal benefits. When mating is not monogamous, high-status males are able to translate status into large reproductive gains, leading to a strong correlation between dominance rank and mating success

across male primates (16). In humans, the reproductive benefits of status reached their peak in premodern states and empires, where sultans, kings, and emperors could control access to a large number of women (17). Studies of the Y chromosome suggest a large increase in male reproductive skew with the rise and spread of agriculture 10,000 y ago (18), and common Y haplotypes can be traced to the lineages of high-status rulers such as Genghis Khan (19, 20). In modern industrial societies with monogamy and low fertility, several studies find that male fertility associates modestly with wealth, largely due to higher childlessness among poorer men (21–24).

Most of human history transpired in small-scale societies, who relied on foraging for subsistence. Observation and archaeology of foragers reveals tremendous variation in status hierarchy (25). In low-density relatively nomadic forager societies, decision-making is typically consensus based (at least among adult men), and status inequality is limited by fluid group membership, coalitional checks on would-be dominants, and cooperative production and interdependence (14, 25–27). Leadership tends to arise occasionally to meet situational demands and typically involves little or no material benefit relative to followers (28, 29). Variance in male reproduction can be small and not appreciably greater than variance in female reproduction (17). Nonetheless, men's hunting among foragers has been attributed in part to status-seeking, because the most successful male hunters attract more sharing partners, more allies in the context of male–male competition, and more fertile, productive wives and extramarital partners (30–33). Some Australian foragers were gerontocratic and maintained high levels of polygyny, even at relatively low population density (34). A minority of known forager societies

Significance

Much of human behavior results from a desire for social status. From an evolutionary perspective, answering the question of why we pursue status must consider how status affects reproduction, especially in nonindustrial societies with natural fertility. In a metaanalysis of 288 results from 33 nonindustrial populations, we find that status is significantly associated with men's reproductive success, consistent with an evolved basis for status pursuit. Status hierarchies have changed dramatically throughout human history, yet we find that the association between status and reproductive success does not depend on subsistence category (foraging, horticulture, pastoralism, agriculture) or how status is measured. These findings suggest no significant increase in selection on status-enhancing traits with the domestication of plants and animals.

Author contributions: C.R.v.R. and A.V.J. designed research, performed research, analyzed data, and wrote the paper.

The authors declare no conflict of interest.

This article is a PNAS Direct Submission.

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This article contains supporting information online at www.pnas.org/lookup/suppl/doi:10.1073/pnas.1606800113/-DCSupplemental.

were sedentary and lived at relatively high population density, including groups from the west coast of North America, who recognized chiefs and who were more polygynous (35, 36). At least in foragers from western North America, polygyny was associated with the privatization of resource extraction sites, such as riverine fish runs (37). In aggregate, however, foragers from the ethnographic and archaeological record may lack the inequality in material resources that contributes to status hierarchy and reproductive inequality in pastoral or agricultural societies (38–40). Furthermore, foragers often lack the population density and associated collective action problems that can cause the rank and file to prefer more institutionalized, even coercive, political leadership (41–44).

The egalitarianism described of many if not most foragers in the ethnographic and archaeological record suggests that status may have been a relatively weak target of selection throughout much of the evolutionary history of modern humans. Positive selection on status may have increased significantly with domestication of plants and animals, due to greater material wealth inequality, institutionalized leadership, and male reproductive skew (17, 18, 39). This “egalitarianism hypothesis” can be tested with quantitative estimates of the relationship between male status and reproductive success (RS) in contemporary foragers compared with nonforagers.

How status translates into RS bears on debates about the evolution of men’s reproductive strategies, particularly men’s parenting vs. mating effort (45, 46). Men’s status pursuit has been framed in terms of mating effort (30), but the deference and admiration of group members can include aid for men’s families that functions as indirect parental investment (13). Positive associations of status with fertility are consistent with mating effort, whereas negative associations with offspring mortality are consistent with parenting effort. If status associates with fertility but not with offspring mortality, this does not necessarily mean that men’s status goals do not concern parenting effort, because there are many extrinsic reasons for offspring mortality. Rather, such results would suggest that by acquiring status, men maximize mating effort more than parenting effort—the “mating effort hypothesis.”

We test the egalitarianism and mating effort hypotheses via a phylogenetic multilevel metaanalysis of 288 statistical associations between male status and RS, from 46 studies in 33 non-industrial societies. We focused on nonindustrial societies to limit effects of widespread contraception, modern medicine, and low fertility norms associated with the demographic transition. We coded all societies by their main mode of subsistence: foraging (subsistence on undomesticated plants and animals), horticulture (cultivation of domesticated plants in garden plots, based on simple tools), pastoralism (heavy reliance on herding domesticated animals), and agriculture (cultivation of domesticated plants using technologies such as plows and traction animals). Different types of status measures were categorized as formidability (e.g., height, weight, strength, warriorship), hunting skill (e.g., return rate, skill ranking), material wealth (e.g., total value of owned goods, income, livestock, or land ownership), and political influence (e.g., headman, influence ranking). Of the status measures, political influence is closest to our definition of status. The other measures better represent traits that tend to confer status, based on their contribution to group-wide judgments of who is dominant or prestigious. Finally, measures of RS were classified as surviving offspring (number of offspring surviving to a certain age), fertility (total offspring born), offspring mortality (e.g., offspring mortality rate, proportion of offspring dying), mating success (e.g., number of wives or affairs, age at marriage or probability of marriage), and wife quality (e.g., wife’s age or interbirth interval, wife’s productivity). All societies were coded by presence or absence of polygyny. The complete dataset is reported in [Dataset S1](#).

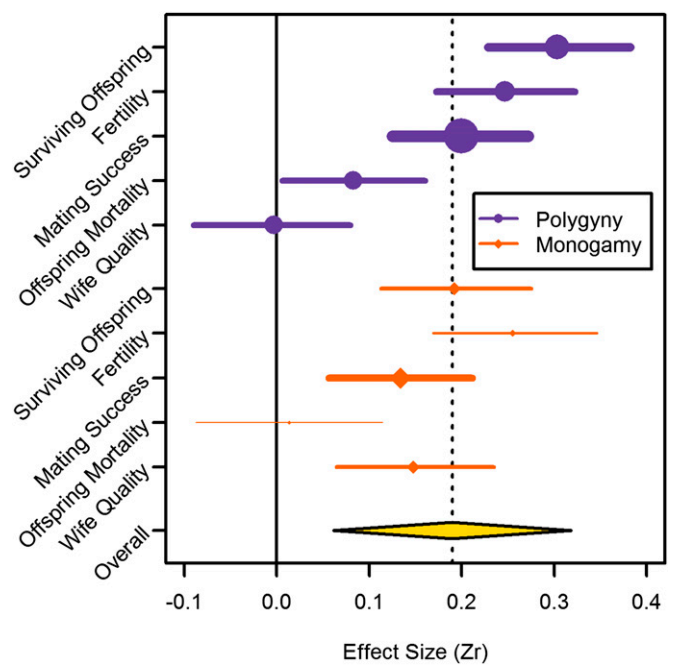


Fig. 1. Variation in the weighted effect of male status on RS, as a function of marriage system (polygyny vs. monogamy) based on averaged coefficients (Table 2). Overall effect size based on intercept-only model. Before metaanalysis, all effect sizes were coded such that positive signs indicate positive contributions to RS (e.g., a negative effect of status on offspring mortality was coded as positive). Point size and line width are proportional to the number of results contributing to each weighted effect size.

Results

We begin with an intercept-only model to estimate the weighted overall effect size for the association between male status and RS. This effect size was significantly greater than zero [$Zr = 0.19$, 95% confidence interval (CI) = 0.09–0.31, $k = 288$; Fig. 1]. Adding a dummy variable for whether the original analysis controlled for age did not change these results. To best explain variation in the effect of status on RS as a function of different covariates, we applied model comparison based on the deviance information criterion (DIC; *Methods*). Table 1 lists the six top-ranked models, with all others receiving DIC weights <0.001. The top two models, together receiving ~90% of DIC weight, contain polygyny and measures of RS and their interaction, but not subsistence type or status type, which only appear in models with low DIC weight. Table 2 presents a summary model for which coefficients were averaged according to DIC weight. In this averaged model, effect sizes vary only minimally as a function of subsistence type or status type, lending no support to the egalitarianism hypothesis. However, there is considerable variation according to RS measure and marriage system (polygyny vs. monogamy), as shown in Fig. 1. In particular, surviving offspring, fertility, and mating success tended to be more strongly associated with status than wife quality and offspring mortality, especially in polygynous groups, thus supporting the mating effort hypothesis. Additionally, the effects of status on wife quality and offspring mortality varied between polygynous and monogamous societies, with wife quality being significantly associated in the latter but not the former, whereas the opposite was true of offspring mortality.

Because polygyny is somewhat confounded with subsistence type, being more common among pastoralists and horticulturalists, the strong polygyny effect found in the model comparison could eclipse any effect of subsistence type. To provide another more direct test of the egalitarianism hypothesis, we therefore fit a model with subsistence type as the only covariate (DIC = –221), which showed no significant differences in the effects of

Table 1. Model comparison showing six top-ranked models explaining variation in the association between male status and reproductive success

Rank	Age	P	RS	St	Su	P × RS	DIC	Weight
1							−320	0.664
2							−318	0.245
3							−314	0.045
4							−314	0.037
5							−310	<0.01
6							−310	<0.01

Cells in gray indicate that a covariate was included in a given model. Polygyny (P), RS type (RS), and their interaction were included in all top-ranked models, whereas status type (St) and subsistence type (Su) only appeared in models with low DIC weight (<0.05). The interaction between subsistence type and RS type did not appear in any model.

status on RS across foragers ($Zr = 0.18$, 95% CI = 0.05–0.32, $k = 64$), horticulturalists ($Zr = 0.17$, 95% CI = 0.03–0.31, $k = 108$), agriculturalists ($Zr = 0.18$, 95% CI = 0.05–0.34, $k = 101$), and pastoralists ($Zr = 0.29$, 95% CI = 0.10–0.44, $k = 15$), though it tended to be slightly higher among the latter (Fig. 2).

To put these results into a larger comparative context, we conducted a phylogenetic multilevel metaanalysis of the association between male rank and mating success among nonhuman primates, focusing on just one dataset (16), which resulted in an overall weighted effect size that was significantly higher than the one for humans ($Zr = 0.80$, 95% CI = 0.43–1.19, $k = 53$; Fig. 2). These results did not change qualitatively when excluding adolescent males from the dataset (16).

No significant publication bias was detected in any model (Methods and Fig. S1), and variation explained at the level of studies and populations, as well as by phylogeny, was minimal throughout (<1% of total variance; Fig. S2).

Discussion

Male social status associates with RS across nonindustrial human societies (Fig. 1). Contra the egalitarianism hypothesis, the effect of status on men's RS does not differ significantly by subsistence type (Fig. 2), despite subsistence-associated variation in political egalitarianism. These results suggest positive selection for traits that

facilitate status acquisition (including men's motivation to seek it) did not increase substantially when foragers began domesticating plants and animals. In fact, the association between status and reproduction during the Pleistocene and Early Holocene may have been greater than we estimate for modern foragers, if modern foragers are more egalitarian due to confinement to marginal habitat (but see ref. 47). Nonetheless, selection for status-enhancing traits throughout human evolution does not entail that the lineages of particular high-status men consistently experienced greater RS (48). The genotypes of high-status men may at times represent fitness peaks, which mutation and sexual recombination break down in successive generations. Status acquisition may depend on adaptations that condition behavior to uncorrelated genetic variation, such as genetic variation that is associated with intelligence or strength and maintained from generation to generation by mutation–selection balance (49).

Why do we not find a stronger effect of status in nonforaging societies? First, effects of status vary considerably within subsistence categories. Yanomamö warriors but not Waorani warriors produced more surviving offspring, despite similar political ecologies and horticultural subsistence (50, 51). In studies of rural communities in 19th-century Finland and Sweden, wealth disparity was low and not a strong predictor of selection in the former (52), but land ownership associated with number of surviving offspring in the latter (53). Even within the same population, the effect of status can vary significantly (Fig. S2). Among the Tsimane' of Amazonian Bolivia, villages differ in political inequality (44) and whether political influence associates with RS (Dataset S1), in proportion to their distance from the market town. There may be more variation within than between subsistence categories in the socioecological factors that favor inequality, including constraints on migration (54), access to and inheritance of monopolizable material wealth (37–40, 55), and collective action problems that catalyze more centralized or coercive leadership (41–44). Future studies of status and reproduction within and across societies should apply more direct metrics of these factors.

Second, the effect of status on men's RS does not vary by status measure, whether status is represented by physical formidability, hunting skill, material wealth, or political influence; this suggests that different, population-specific means of gaining status can be equally effective when it comes to men's RS. William Irons, one of the first researchers to test the status–reproduction relationship in a small-scale society, made a similar claim: “In

Table 2. Averaged model coefficients based on DIC weights (Table 1)

Coefficient	Posterior mean	Lower 95% CI	Upper 95% CI
Baseline	0.303	0.230	0.375
RS type = fertility	0.247	0.180	0.321
RS type = mating success	0.200	0.128	0.266
RS type = offspring mortality	0.084	0.010	0.158
RS type = wife quality	−0.003	−0.086	0.076
Polygyny = absent	0.190	0.111	0.264
RS type = fertility, polygyny = absent	0.254	0.169	0.342
RS type = mating success, polygyny = absent	0.133	0.063	0.208
RS type = offspring mortality, polygyny = absent	0.012	−0.088	0.105
RS type = wife quality, polygyny = absent	0.146	0.067	0.227
Age control = no	0.302	0.229	0.378
Subsistence = horticulture	0.301	0.230	0.375
Subsistence = pastoralism	0.308	0.234	0.380
Subsistence = agriculture	0.305	0.235	0.380
Status = hunting ability	0.304	0.232	0.377
Status = political influence	0.304	0.232	0.377
Status = wealth	0.304	0.233	0.378

The baseline level refers to RS type = surviving offspring, subsistence = foraging, status = formidability, polygyny = present, age control = yes, and other coefficients represent deviations from this baseline as indicated. All effect sizes were coded such that positive signs indicate positive contributions to RS (e.g., a negative effect of status on offspring mortality was coded as positive). Though there is substantial variation according to RS type and polygyny (Fig. 1), effect sizes are largely the same for different subsistence types or different status types.

offspring, is echoed by the overall lower effects of status on reproduction in humans compared with nonhuman primates (Fig. 2). This reduction in reproductive payoffs to high status could reflect increased interdependence in food production, greater ability to form leveling coalitions, increased payoffs to mate-guarding due to changes in sex ratio, or increased opportunities for female choice (75, 78). Arranged marriage has likely been common since at least the early migrations of humans out of Africa (79), but even when male kin control marriage arrangements, females may exert significant mate choice via surreptitious extrapair copulation (80).

Why women pursue status may differ from men. Some have argued that sexual selection theory predicts men are more driven to compete for privileged positions within community-wide coalitions, whereas women are more motivated by a strong dyadic support network (81, 82). Women also face constraints on community-wide networking due to the sexual division of labor. However, women in nonindustrial societies often have important roles in community decision-making (81, 83) and conflict resolution (84, 85), particularly where inheritance is matrilineal (81). Better understanding of how women's status is shaped by socioecological variation may be a boon for policies aimed at increasing the representation of women in leadership positions, and remains an open field of investigation for evolutionary anthropologists.

Methods

A multifaceted search produced the 46 studies that comprise the meta-analysis. Journals that focus on anthropology (e.g., *Current Anthropology*) and the behavioral sciences (e.g., *Evolution and Human Behavior*) as well as more general science journals (e.g., *Science*) were repeatedly searched for terms related to social status, including status, hunting, leadership, and wealth. The reference list of each relevant study was scoured for additional studies. Criteria for inclusion were a nonindustrial sample, at least one measure of male social status, and at least one quantitative association of male social status with a measure of RS. We excluded two studies that restricted their analysis to status grades within the nobility rather than across the entire population (86, 87). Status measures could include direct measures of status (e.g., political leadership) or more indirect measures of status (e.g., wealth) that the study authors associated with status in their sample. Measures of RS could include number of offspring surviving to a certain age, fertility, offspring mortality rate, and their proximate determinants, including number of mates, mates' interbirth interval, the health of mates and offspring, mates' productivity, and age of spouses at marriage.

Effect sizes were calculated from the published information on statistical tests (e.g., r , d , or F statistics, sample size n , and/or associated P values), and all effect sizes were transformed into Fisher's Z (Z_r) and its associated variance following standard metaanalytic procedures (88). All effect sizes were coded such that positive signs indicate positive contributions to RS (e.g., a negative effect of status on offspring mortality or age of marriage was coded as positive).

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Where incomplete or insufficient information was published [e.g., $P < 0.05$, not significant (n.s.)], we solicited complete information or raw data from the first author for studies published in the past 5 y, and otherwise conservatively used the upper boundary of the published interval ($P < 0.05$) or assumed an effect size of zero (n.s.). A dummy variable was included to indicate whether the analysis controlled for men's age, either by focusing only on completed fertility, by including age as a covariate, by analyzing different age cohorts separately, or by age-matching high- and low-status men. The collection of all analyzed data was approved by the universities affiliated with the respective studies.

We used Bayesian phylogenetic multilevel metaregression (89, 90) to estimate an overall weighted effect size and to model the influence of different covariates. We controlled for cultural and phylogenetic history by including a phylogenetic tree (Fig. S2) based on a recent supertree of human cultures that combines linguistic and genetic information (91). See *SI Methods* for details. Each effect size was weighted by its SE. To account for repeated measures, variation at the level of studies and populations was modeled using random effects. For our model comparison, the full model considered was (see model 11 in ref. 89):

$$\begin{aligned} Z_r \sim & \text{age control}(\text{binary}) + \text{status type} + \text{subsistence type} + \text{RS measure} \\ & + \text{polygyny}(\text{binary}) + \text{subsistence} \times \text{RS} + \text{polygyny} \times \text{RS} \\ & + \text{random effects for study and population} + \text{phylogenetic signal} \\ & + \text{measurement errors} + \text{residuals} \end{aligned}$$

All possible combinations of these covariates were fit and ranked based on their DIC weight using the MuMIn package (92). Coefficients were then averaged according to DIC weight to produce an average model, setting coefficients of parameters not included in a model to zero (93). In all analyses the proportion of total variance explained by the study-level and population-level random effects as well as the phylogenetic signal (90) was <0.01 , indicating that behavioral variation within levels far outweighed variation between levels (Fig. S2). We tested for publication bias using Egger's regression on the metaregression residuals (90); no analysis revealed significant publication bias (Fig. S1), consistent with negative results in this field being published in high-impact journals (51). For the analysis of the nonhuman primate sample, a consensus tree was downloaded from the 10kTrees website (94) and intraspecific variation was modeled using a species-level random effect (90).

All analyses were run in R 3.0.2 (95) using the MCMCglmm package (89) with standard inverse gamma priors for random effects, a burn-in period of 20,000 iterations, and 100,000 iterations in total thinned to 8,000 samples. Model convergence was assessed visually by plotting time series and histograms of the Markov chains (Fig. S3) and by calculating the Gelman–Rubin diagnostic on multiple runs of the same model (<1.1 in all analyses).

ACKNOWLEDGMENTS. We thank Michael Gurven, Paul Hooper, and three anonymous reviewers who provided comments on earlier drafts that led to substantial improvements; Shinichi Nakagawa for sharing his R code to run Egger's regression; and Eric Alden Smith, John Ziker, and Luke Glowacki who generously provided analyses or data associated with their published studies.

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